Extending univariate methods

Applied Multivariate Statistics – Spring 2013



Overview

- Multivariate t-test (one sample, two samples)
- MANOVA
- Multivariate Linear Regression

Revision: One-sample z-Test

- 1. Model: $X_1, ..., X_n \sim N(\mu, \sigma_X^2)$ iid, σ_X known
- 2. Hypotheses: $H_0: \mu = \mu_0, H_A: \mu \neq \mu_0$
- 3. Test statistics:

$$T = \frac{\overline{X}_n - \mu_0}{\sigma_{\overline{X}_n}} = \sqrt{n} \frac{\overline{X}_n - \mu_0}{\sigma_X}$$

If H_0 is true: $\overline{X}_n \sim N(\mu_0, \frac{\sigma_X^2}{n})$ and thus $T \sim N(0,1)$

- 4. Make observation of test statistics: t
- 5. Compute p-value: Probability of seeing something as extreme as t or even more extreme than t if H_0 is true: P(|T| > |t|)

Revision: One-sample t-Test

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$$T = \frac{\overline{X}_n - \mu_0}{\hat{\sigma}_{\overline{X}_n}} = \sqrt{n} \frac{\overline{X}_n - \mu_0}{\hat{\sigma}_X}$$

If H_0 is true: $\overline{X}_n \sim N(\mu_0, \frac{\sigma_X^2}{n})$ and thus $T \sim t_{n-1}$

- 4. Make observation of test statistics: t
- 5. Compute p-value: Probability of seeing something as extreme as t or even more extreme than t if H_0 is true: P(|T| > |t|)

Hotelling's one-sample T-Test: Σ known

- 1. Model: $X_1, ..., X_n \sim N(\mu, \Sigma)$ *iid*, Σ known; p dimensions
- 2. Hypotheses: $H_0: \mu = \mu_0, H_A: \mu \neq \mu_0$
- 3. Test statistics:

$$T = n \left(\overline{X_n - \mu_0} \right)^T \Sigma^{-1} \left(\overline{X_n - \mu_0} \right) \leftarrow$$

If H_0 is true: $T \sim \chi_p^2$

Squared Mahalanobis Distance between sample mean and
$$\mu_0$$

- 4. Make observation of test statistics: t
- 5. Compute p-value: Probability of seeing something as extreme as t or even more extreme than t if H_0 is true: P(|T| > |t|)

Hotelling's one-sample T-Test: Σ unknown

- 1. Model: $X_1, ..., X_n \sim N(\mu, \Sigma)$ *iid*, Σ unknown; p dimensions
- 2. Hypotheses: $H_0: \mu = \mu_0, H_A: \mu \neq \mu_0$
- 3. Test statistics:

$$T = n \left(\overline{X_n - \mu_0} \right)^T S^{-1} \left(\overline{X_n - \mu_0} \right) \leftarrow$$

If H_0 is true: $T \sim F_{p,n-p}$

Estimated Sq. Mahalanobis Distance between sample mean and
$$\mu_0$$

- 4. Make observation of test statistics: t
- 5. Compute p-value: Probability of seeing something as extreme as t or even more extreme than t if H_0 is true: P(|T| > |t|)

R: Function "HotellingsT2" in package "ICSNP"

F distribution



Example: Change in Pulmonary Response after Exposure to Cotton Dust



Revision: Two-sample t-Test

1. Model:
$$X_1, \dots, X_n \sim N(\mu_X, \sigma_X^2)$$
 iid, σ_X unknown
 $Y_1, \dots, Y_m \sim N(\mu_Y, \sigma_X^2)$ iid Can

2. Hypotheses: $H_0: \mu_X = \mu_Y, H_A: \mu_X \neq \mu_Y$

Can be extended to $\sigma_X \neq \sigma_Y$

3. Test statistics:

$$T = \frac{(\overline{X}_n - \overline{Y}_n) - (\mu_X - \mu_Y)}{\hat{\sigma}_{\overline{X}_n - \overline{Y}_n}}$$

If H_0 is true: $T \sim t_{n+m-2}$

- 4. Make observation of test statistics: t
- 5. Compute p-value: Probability of seeing something as extreme as t or even more extreme than t if H_0 is true: P(|T| > |t|)

Hotelling's Two-Sample T-Test: Σ unkown, but equal in both groups

- 1. Model: $X_1, ..., X_n \sim MVN(\mu_X, \Sigma)$ *iid*, Σ unknown, p dims. $Y_1, ..., Y_m \sim MVN(\mu_Y, \Sigma)$ *iid*
- 2. Hypotheses: $H_0: \mu_X = \mu_Y, H_A: \mu_X \neq \mu_Y$
- 3. Test statistics:

$$T = \frac{(n+m-p-1)nm}{(n+m)(n+m-2)p} (\overline{X}_n - \overline{Y}_n)^T S^{-1} (\overline{X}_n - \overline{Y}_n)$$

If H_0 is true: $T \sim F_{p,n+m-p-1}$

- 4. Make observation of test statistics: t
- 5. Compute p-value: Probability of seeing something as extreme as t or even more extreme than t if H_0 is true: P(|T| > |t|)
 - R: Function "HotellingsT2" in package "ICSNP"

Example: Quality control for screws



20 screws:

- winding [mm]
- length [mm]
- diameter [mm]

Plant struck by lightning: Machines still adjusted correctly?

- 15 screws:
- winding
- length
- diameter

Revision: One-way ANOVA

Are the expected values in three groups the same?



ANOVA:

- Compare variation within groups and between groups
- Assume normality

⇒ p-Values can be computed

MANOVA

Are the multi-dimensional expected values in three groups the same?





• MANOVA:

- Compare within groups and between groups covariance matrices (test statistics based on eigenvalues)

- Assume normality
 Wilks test: p-Values can be computed
- R: Function "manova" and "summary(..., test = "Wilks")"

Revision: Univariate (Multiple) Linear Regression

- N samples, p predictors, 1 response
- Univariate Linear Regression model:

$$Y = \beta_0 + \sum_{j=1}^p X_j \beta_j + \epsilon = f(X) + \epsilon$$

For N samples using matrix notation:

$$Y = X\beta + E$$

where
Y: N*1 matrix, X: N*(p+1), β : (p+1)*1, E: N*7

- Criterion to optimize: $RSS(\beta) = \sum_{i=1}^{N} (y_i f(x_i))^2$
- Solution: $\hat{\beta} = (X^T X)^{-1} X^T Y$

Multivariate (Multiple) Linear Regression

- N samples, p predictors, K responses
- Univariate Linear Regression model for each response:

 $Y_k = \beta_{0k} + \sum_{j=1}^p X_j \beta_{jk} + \epsilon_k = f_k(X) + \epsilon_k$

 $Cov(\epsilon) = \Sigma$, errors between responses can be correlated For N samples using matrix notation:

Y = XB + E

where

Y: N*K matrix, X: N*(p+1), B: (p+1)*K, E: N*K

- Criterion to optimize: $RSS(B; \Sigma) = \sum_{i=1}^{N} (y_i f(x_i))^T \Sigma^{-1} (y_i f(x_i))$
- Solution: $\hat{B} = (X^T X)^{-1} X^T Y$
- Surprising result: Estimates and even confidence intervals are the same if doing K univariate multiple regressions!

Is MANOVA and Multivariate Linear Regression useful?

- Multivariate Regression, MANOVA not well supported in statistical software (including R)
- Useful, if you want to test if a predictor has an influence on any response
- Possible in theory, but not well supported:
 - simultaneous confidence intervals for several parameters
 - Tests among parameters of different responses
- R: Function "Im" with matrix as y and "summary(..., test = "Wilks")"

Concepts to know

- Hotelling's T-test
- Idea of MANOVA and Multivariate Regression

R functions to know

- "HotellingsT2"
- "Manova"
- "Im" with y being a matrix