

## Series 6

1. The dataset `heart.dat` contains data for 99 people sorted by age. In each age group the total number of individuals ( $m_i$ ) is known, as well the number of those with symptoms of heart disease ( $N_i$ ). The goal of this exercise is to estimate the probability of having such symptoms as a function of age using logistic regression.

The data is located at <http://stat.ethz.ch/Teaching/Datasets/heart.dat>.

- a) In contrast to the binary classification example in the lecture notes (page 56), the response variable  $N$  has not a Bernoulli, but a binomial distribution:  $N_1, \dots, N_n$  independent,  $N_i \sim \text{Binomial}(m_i, \pi(x_i))$ .

Show that the log-likelihood is in this case

$$\ell(\beta; (x_1, m_1, N_1), \dots, (x_n, m_n, N_n)) = \sum_{i=1}^n \left[ \log \binom{m_i}{N_i} + N_i g(\beta; x_i) - m_i \log \left( 1 + e^{g(\beta; x_i)} \right) \right],$$

where  $g(\beta; x) = \beta_0 + \beta_1 x$  is the model function for the logistic transform of  $\pi(x)$  (see Formula (6.6) in the lecture notes).

- b) Write an R function `neg.ll(beta, data)` that calculates the *negative* log-likelihood

$$-\ell(\beta; (x_1, m_1, N_1), \dots, (x_n, m_n, N_n))$$

that you derived in task a). `beta` is a vector with two entries  $\beta_0$  and  $\beta_1$ , and `data` is a data frame with columns `age`, `m` and `N` (as in `heart.dat`).

Make a contour plot of the negative log-likelihood of the `heart` dataset in the range  $-10 \leq \beta_0 \leq 10$ ,  $-1 \leq \beta_1 \leq 1$ .

**R hint:** use a command like

```
> contour(beta0.grid, beta1.grid, neg.ll.values)
```

`beta0.grid` and `beta1.grid` are equidistant grids of values of  $\beta_0$  and  $\beta_1$  in the region of interest; use, e.g.

```
> beta0.grid <- seq(-10, 10, length = 51)
```

`neg.ll.values` is a matrix of negative log-likelihood values for the different values of  $\beta_0$  and  $\beta_1$ .

- c) Estimate the parameters  $\beta_0$  and  $\beta_1$  of the model function (see task a)) using the R function `glm`. Does age influence this probability in a significant way? How do you interpret the sign of the coefficient of `age`?

Compare the estimates from `glm` with estimates you get when minimizing the negative log-likelihood function you implemented in task b).

**R hint:** the logistic regression model can be fitted by using the command

```
> fit <- glm(cbind(N, m - N) ~ age, family = binomial, data = heart)
```

Binomial responses  $N_i \sim \text{Bin}(m_i, \pi_i)$  for  $m_i > 1$  should be entered as a (two-column) matrix, with the number of “successes” ( $N_i$ ) in the first column and the number of “failures” ( $m_i - N_i$ ) in the second.

To minimize your function `neg.ll` from task b), use

```
> optim(c(0, 0), neg.ll, data = heart)
```

The first argument is the start value used for numerical optimization.

- d) Plot the probability estimate against age. At what age would you expect 10%, 20%, ..., 90% of people to have symptoms of heart disease? Discuss your results.

**R hint:** you can obtain probability estimates at arbitrary ages `new.age` by using the command

```
> predict(fit, newdata = data.frame(age = new.age), type = "response")
```

2. a) **Quadratic Discriminant Analysis (QDA)**

Assume the normal model  $X|Y = j \sim \mathcal{N}_p(\mu_j, \Sigma_j)$ ,  $\mathbb{P}[Y = j] = p_j$ ,  $\sum_{j=0}^{J-1} p_j = 1$ . Show that (6.2) and (6.4) lead to

$$\hat{\delta}_j^{QDA}(x) = -\log(\det(\hat{\Sigma}_j))/2 - (x - \hat{\mu}_j)^\top \hat{\Sigma}_j^{-1} (x - \hat{\mu}_j)/2 + \log(\hat{p}_j).$$

b) **Linear Discriminant Analysis (LDA)**

Use the result from a) and replace  $\hat{\Sigma}_j$  by  $\hat{\Sigma}$  to get

$$\begin{aligned} \hat{\delta}_j^{LDA}(x) &= x^\top \hat{\Sigma}^{-1} \hat{\mu}_j - \hat{\mu}_j^\top \hat{\Sigma}^{-1} \hat{\mu}_j / 2 + \log(\hat{p}_j) \\ &= (x - \hat{\mu}_j / 2)^\top \hat{\Sigma}^{-1} \hat{\mu}_j + \log(\hat{p}_j). \end{aligned} \quad (1)$$

c) The LDA decision function can be written as (see (1) above)

$$\hat{\delta}_j(x) = x^\top b_j + c_j,$$

where  $b_j \in \mathbb{R}^p$  and  $c_j \in \mathbb{R}$ . Assume that we only have two classes ( $j = 0, 1$ ). Use the equation above to characterize the decision boundary.

d) **Small Simulation**

Use the R-code below to generate data samples from three groups of normal distributions; change the covariance matrix and mean vectors if you like:

```
library(MASS)    ## Needed for lda/qda and mvrnorm
## Read in a function that plots LDA/QDA decision boundaries
source("http://stat.ethz.ch/education/semesters/ss2012/CompStat/predplot.R")
## Covariance Matrix
sigma <- cbind(c(0.5, 0.3), c(0.3, 0.5))
## Mean vectors
mu1 <- c(3, 1.5)
mu2 <- c(4, 4)
mu3 <- c(8.5, 2)
m <- matrix(0, nrow = 300, ncol = 3)
## Grouping vector
m[,3] <- rep(1:3, each = 100)
## Simulate data
m[1:100,1:2] <- mvrnorm(n = 100, mu = mu1, Sigma = sigma)
m[101:200,1:2] <- mvrnorm(n = 100, mu = mu2, Sigma = sigma)
m[201:300,1:2] <- mvrnorm(n = 100, mu = mu3, Sigma = sigma)
m <- data.frame(m)
Perform LDA and plot the results:
fit <- lda(x = m[,1:2], grouping = m[,3])
predplot(fit, m)
```

Manually calculate (see c)) the boundary between group 1 and 2. Add your solution to the plot with `abline()`

**Hint:**

If `A <- fit$scaling`, it holds (in the case of  $p + 1$  groups in  $\mathbb{R}^p$ ) that  $\hat{\Sigma}^{-1} = AA^\top$ . The means and prior probabilities can also be found in the `lda`-object. However, you may also want to do everything on your own, i.e., without using the result of `lda`; in this case, you can use the estimators for  $\hat{\mu}_j$  and  $\hat{\Sigma}$  given in Chapter 6.3.1 of the lecture notes, just above Formula (6.5).

**Preliminary discussion:** Friday, April 20.

**Deadline:** Friday, April 27.