

Extension of the Cox Proportional Hazards Model for Time-Dependent Variables

Tulasi Agnihotram, Gaby Binder, Fabian Frei
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1 Review

Cox PH-Model: $h(t, \mathbf{X}) = h_0(t) \cdot \exp[\sum_{i=1}^p \beta_i X_i]$

Hazard ratio: $\hat{H}R(t) = \frac{\hat{h}(t, X^*(t))}{\hat{h}(t, X(t))} = \exp[\sum_{i=1}^p \hat{\beta}_i [X_i^* - X_i]]$

PH assumption: The hazard ratio is independent of time: $\frac{h(t, X^*)}{h(t, X)} = \theta$

Methods for checking the PH assumption:

- Graphical
- Time-dependent covariates
- Goodness-of-fit test

What can be done if the PH assumption is not met:

- Use a stratified Cox procedure
- Use an extended Cox model

2 Time-dependent Variables

Definition: Any variable whose values differ over time.

Example:

- Race: time-independent
Race \times t: time-dependent

There are 3 types of variables:

- Defined variables
- Internal variables
- Ancillary variables

3 The Extended Cox Model for Time-dependent Variables

$$h(t, \mathbf{X}(t)) = h_0(t) \cdot \exp\left[\sum_{i=1}^{p_1} \beta_i X_i + \sum_{i=1}^{p_2} \delta_i X_i(t)\right]$$

$\mathbf{X}(t) = (X_1, \dots, X_{p_1}, X_1(t), \dots, X_{p_2}(t))$ denotes all predictors
 X_i denotes the i^{th} time-independent variable
 $X_i(t)$ denotes the i^{th} time-dependent variable

- The ML procedure is used to estimate the regression coefficients
- The model assumes that the hazard at time t depends on the value of $X_i(t)$ at the SAME time t
- We can modify the model to allow lag-time

4 The Hazard Ratio for the Extended Cox Model

Extended hazard ratio:

$$\hat{H}R(t) = \frac{\hat{h}(t, X^*(t))}{\hat{h}(t, X(t))} = \exp\left[\sum_{i=1}^{p_1} \hat{\beta}_i [X_i^* - X_i] + \sum_{i=1}^{p_2} \hat{\delta}_i [X_i^*(t) - X_i(t)]\right]$$

The HR depends on time. That's why the PH assumption is not satisfied.

5 Assessing Time-independent Variables that do not satisfy the PH Assumption

General formula for assessing the PH assumption:

$$h(t, \mathbf{X}(t)) = h_0(t) \cdot \exp\left[\sum_{i=1}^{p_1} \beta_i X_i + \sum_{i=1}^{p_2} \delta_i X_i g_i(t)\right]$$

Here the $g_i(t)$ is a function of time. It is important which form we choose for $g_i(t)$ in the model.

We do a test for assessing the PH assumption using the Likelihood ratio test

$$LR = -2 \log(L_{PH \text{ model}}) - (-2 \log(L_{\text{ext. Cox model}})) \sim \chi_p^2 \text{ under } H_0$$

Nullhypothesis H_0 : $\delta_1 = \dots = \delta_p = 0$

If the test is significant then the extended Cox model is preferred.

We can choose $g_i(t)$ as a heaviside function $g_i(t) = \begin{cases} 1 & \text{if } t \geq t_0 \\ 0 & \text{if } t < t_0 \end{cases}$

In a model, we can use one or more heaviside functions. If we use heaviside functions, the HR yields constant values for different time intervals.

6 An Application of the Extended Cox Model: Treatment of Heroin Addiction

We compare two methadone maintenance clinics. Clinic 2 has always higher retention probabilities than clinic 1. The difference is very significant after one year of treatment. Because the two curves in the $-\ln(-\ln(S))$ plot are not parallel, the variable clinic doesn't satisfy the PH assumption. Two extended Cox models were considered:

- Heaviside functions to obtain two distinct hazard ratios. One for less than one year and the other for greater than one year
- A time-dependent variable that allows for the two survival curves to diverge over time

7 The Extended Cox Likelihood

Example:

	TIME	STATUS	SMOKE
Barry	2	1	1
Gary	3	1	0
Harry	5	0	0
Larry	8	1	1

Cox PH model: $h(t) = h_0(t)e^{\beta_1 SMOKE}$

Cox Likelihood:

$$L = \left[\frac{h_0(2)e^{\beta_1}}{h_0(2)e^{\beta_1} + h_0(2)e^0 + h_0(2)e^0 + h_0(2)e^{\beta_1}} \right] \\ \times \left[\frac{h_0(3)e^0}{h_0(3)e^0 + h_0(3)e^0 + h_0(3)e^{\beta_1}} \right] \\ \times \left[\frac{h_0(8)e^{\beta_1}}{h_0(8)e^{\beta_1}} \right]$$

Cox extended model: $h(t) = h_0(t)e^{\beta_1 SMOKE + \beta_2 SMOKE \times TIME}$

Extended Cox Likelihood:

$$L = \left[\frac{h_0(2)e^{\beta_1 + 2\beta_2}}{h_0(2)e^{\beta_1 + 2\beta_2} + h_0(2)e^0 + h_0(2)e^0 + h_0(2)e^{\beta_1 + 2\beta_2}} \right] \\ \times \left[\frac{h_0(3)e^0}{h_0(3)e^0 + h_0(3)e^0 + h_0(3)e^{\beta_1 + 3\beta_2}} \right] \\ \times \left[\frac{h_0(8)e^{\beta_1 + 8\beta_2}}{h_0(8)e^{\beta_1 + 8\beta_2}} \right]$$

References

1. Kleinbaum, D.G. and Klein, M. (2005). *Survival Analysis. A self-learning text*. Springer.
2. <ftp://stat.ethz.ch/WBL/Source-WBL-2/R/TK.R.functions.R>
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