



Split Plot Designs

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What is a Split Plot Design? (Oehlert, 2000, Chapter 16.1)

- A split plot design is a **special case** of a **factorial treatment structure**.
- It is used when some factors are **harder** (or more expensive) to **vary** than others.
- Basically a split plot design consists of two experiments with **different experimental units** of different “size”.
- E.g., in agronomic field trials certain factors require “large” experimental units, whereas other factors can be easily applied to “smaller” plots of land.
- Let us have a look at an example...

Example I: Irrigation and Corn Variety (Oehlert, 2000)

- Consider the following factorial problem:
 - 3 different **irrigation levels**
 - 4 different **corn varieties**
 - Response: **biomass**
 - Available resources: **6 plots of land**
- By definition we **can not** vary the irrigation level on a too small scale.
- We are “forced” to use “large” experimental units for the irrigation level factor.
- Assume that we can use a specific irrigation level on each of the 6 plots.



Example I: Irrigation and Corn Variety

- Randomly assign each irrigation level to 2 of the plots (the so called **whole plots** or **main plots**).
- In every of the plots, randomly assign the 4 different corn varieties to the so called **split plots**.

| | | | | | |
|---|---|---|---|---|---|
| 4 | 4 | 1 | 3 | 2 | 2 |
| 2 | 1 | 4 | 1 | 4 | 1 |
| 3 | 3 | 2 | 2 | 1 | 4 |
| 1 | 2 | 3 | 4 | 3 | 3 |

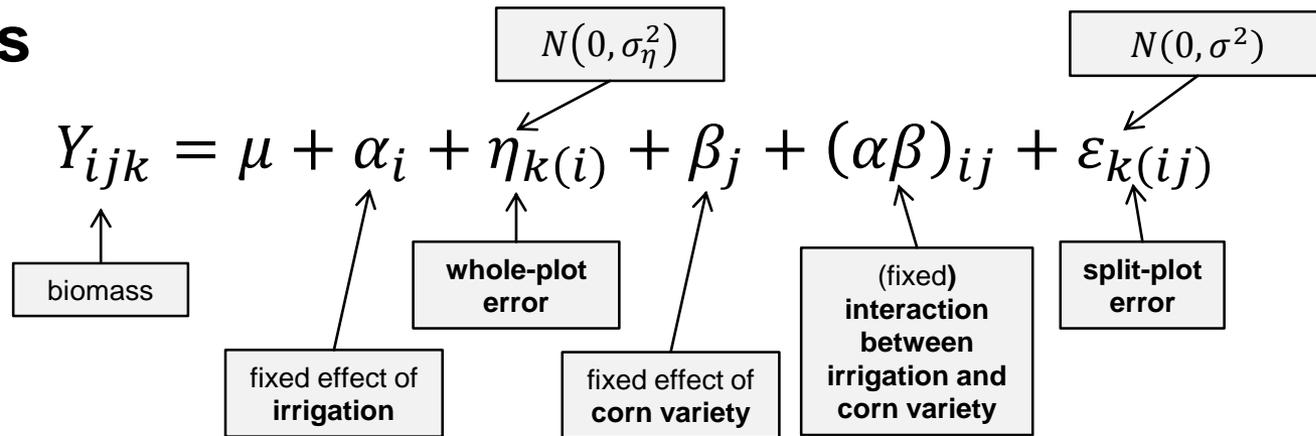
- **Two independent randomizations** are being performed!
- We also call irrigation level the **whole-plot factor** and corn variety the **split-plot factor**.

Example I: Irrigation and Corn Variety

- **Whole plots** (plots of land) are the experimental units for the **whole-plot factor** (irrigation level).
- **Split plots** (subplots of land) are the experimental units for the **split-plot factor**.
- In the split-plot “world”, whole plots act as **blocks**.
- Basically, we are performing **two different experiments in one**:
 - each experiment has its own **randomization**
 - each experiment has its own idea of **experimental unit**

Example I: Irrigation and Corn Variety

- How can we model such kind of data?
- We use a **mixed model** formulation with two **different errors**



- This means: Observations in the same whole plot share the same whole-plot error $\eta_{k(i)}$.
- In R, this model is easily fitted using `lmer` with a **random effect** (better terminology: **error**) of the form `(1 | whole.plot)`

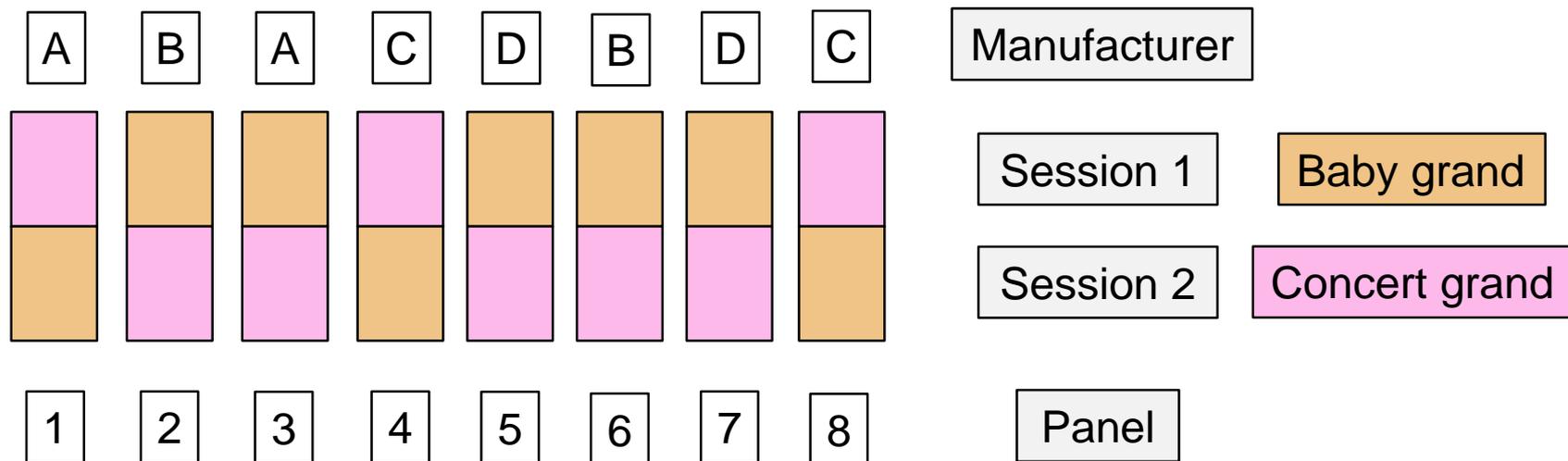
Example II: Pianos (Oehlert, 2000)



- **Two piano types** (baby grand / concert grand) from each of **4 manufacturers**.
- 40 music students are divided at random into **8 groups** (“**panels**”) of 5 students each.
- Two panels are assigned at random to each manufacturer (= 2 panels per manufacturer).
- Each panel goes to the concert hall and hears (blindfolded) the sound of **both pianos** (in random order).
- Response: Average rating of the 5 students in the panel (hence, student is “only” measurement unit here).

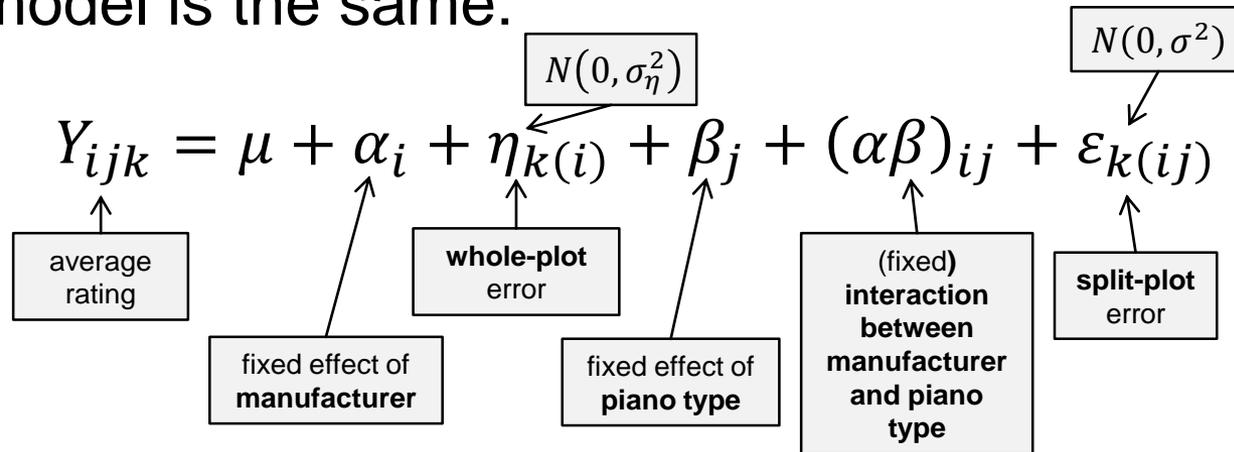
Example II: Pianos

- The **whole plots** are the 8 **panels**.
- The **whole-plot factor** is the **manufacturer**.
- The **split plots** are the **two sessions**.
- The **split-plot factor** is the **piano type** (baby vs. concert grand).



Example II: Pianos

- The model is the same:



- Again: This means that observations in the same whole-plot share the same whole-plot error $\eta_{k(i)}$ and are therefore **not independent**.

Example III: Oats



- Dataset `oats` from R-package `MASS`.
- As stated in the help file:

*The yield of oats from a **split-plot field trial** using three varieties and four levels of manurial treatment. The experiment was laid out in 6 blocks of 3 main plots, each split into 4 sub-plots. The varieties were applied to the main plots and the manurial treatments to the sub-plots.*
- Overview of data:
 - 6 different **blocks (B)**
 - 3 different **varieties (V)**
 - 4 different **nitrogen treatments (N)**
 - **Response (Y):** Yields (in $\frac{1}{4}$ lbs per sub-plot, each of area $\frac{1}{80}$ acre).
- Let us first have a **graphical overview** of the experimental design.

Example III: Oats

| | | |
|---|---|---|
| 4 | 4 | 1 |
| 2 | 1 | 3 |
| 3 | 3 | 2 |
| 1 | 2 | 4 |

I

| | | |
|---|---|---|
| 2 | 1 | 1 |
| 1 | 2 | 4 |
| 3 | 4 | 2 |
| 4 | 3 | 3 |

II

| | | |
|---|---|---|
| 3 | 3 | 2 |
| 2 | 2 | 3 |
| 1 | 4 | 4 |
| 4 | 1 | 1 |

III

| | | |
|---|---|---|
| 1 | 1 | 3 |
| 2 | 3 | 2 |
| 4 | 2 | 1 |
| 3 | 4 | 4 |

IV

| | | |
|---|---|---|
| 3 | 4 | 3 |
| 2 | 1 | 4 |
| 4 | 2 | 1 |
| 1 | 3 | 2 |

V

| | | |
|---|---|---|
| 2 | 3 | 1 |
| 1 | 4 | 4 |
| 4 | 2 | 2 |
| 3 | 1 | 3 |

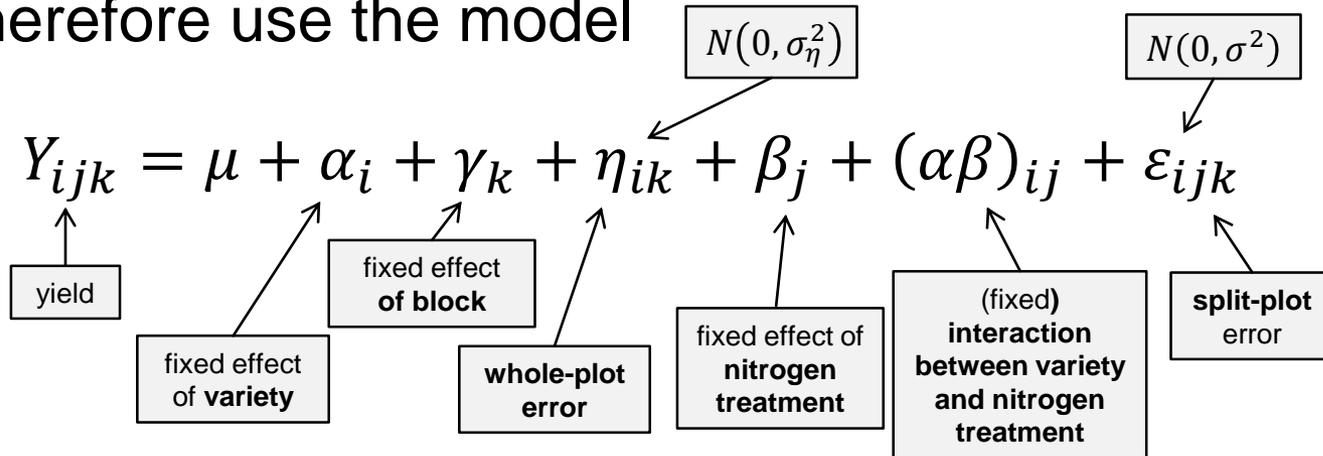
VI

Example III: Oats

- This is a more complicated design as before as we have an additional **block factor**.
- A **whole-plot** is given by a plot of land in a block.
- The **whole-plot factor** is variety.
- A **block design (RCB)** was used at the whole-plot level.
- A **split plot** is given by a **subplot of land**.
- The **split-plot factor** is given by **nitrogen treatment**.

Example III: Oats

- We have an **RCB** for the whole-plot factor.
- The **experimental unit** on the whole-plot level is given by the **combination of block and variety**.
- We therefore use the model



Example III: Oats

- In R we use the `lmer` function with an extra random effect (error) per **combination** of **block** and **variety**.

- We get the following output

```
> fit.lme <- lmer(Y ~ B + V * N + (1 | B:V), data = oats)
> anova(fit.lme)
Analysis of Variance Table of type III with Satterthwaite
approximation for degrees of freedom
```

| | Sum Sq | Mean Sq | NumDF | DenDF | F.value | Pr(>F) |
|-----|---------|---------|-------|-------|---------|---------------|
| B | 4675.0 | 935.0 | 5 | 10 | 5.280 | 0.01244 * |
| V | 526.1 | 263.0 | 2 | 10 | 1.485 | 0.27239 |
| N | 20020.5 | 6673.5 | 3 | 45 | 37.686 | 2.458e-12 *** |
| V:N | 321.8 | 53.6 | 6 | 45 | 0.303 | 0.93220 |

- Observe that the test for variety uses 2 and 10 degrees of freedom.
- Why? Let us have a closer look at the potential ANOVA table on the whole-plot level.

Example III: Oats

- On the whole-plot level we have the following ANOVA table:

| <i>Source</i> | <i>df</i> |
|--------------------|---------------|
| Block | 5 |
| Variety | 2 |
| Error (whole-plot) | 10 (= 17 – 7) |
| Total | 17 (= 18 – 1) |

- Think of averaging “away” the nitrogen factor, hence we have **one observation per combination of block and variety**.
- Technically speaking, variety is tested against the **interaction** of block and variety.

Example III: Oats

- This also reveals a problem: We don't have too many error df's left to test the whole-plot factor (only 10).
- In contrast, we test everything involving the **split-plot** factor against the **residual error**, which has 45 df's.

- Remember:

```
> fit.lme <- lmer(Y ~ B + V * N + (1 | B:V), data = oats)
> anova(fit.lme)
```

Analysis of Variance Table of type III with Satterthwaite approximation for degrees of freedom

| | Sum Sq | Mean Sq | NumDF | DenDF | F.value | Pr(>F) | |
|-----|---------|---------|-------|-------|---------|-----------|-----|
| B | 4675.0 | 935.0 | 5 | 10 | 5.280 | 0.01244 | * |
| V | 526.1 | 263.0 | 2 | 10 | 1.485 | 0.27239 | |
| N | 20020.5 | 6673.5 | 3 | 45 | 37.686 | 2.458e-12 | *** |
| V:N | 321.8 | 53.6 | 6 | 45 | 0.303 | 0.93220 | |

- Hence, all effects involving the whole-plot factor are estimated **less precisely** and tests are **less powerful**.

General Situation

- Split-plot designs can also arise in (much) more complicated designs.
- There can be more than one whole-plot factor. E.g., think of a two-way factorial on the whole-plot level.
- In addition, there can be more than one factor on the split-plot level.
- To get the correct model we “only” have to **follow “the path of randomization”**.
- For every “level” (whole-plot / split-plot) of the experiment we have to introduce a corresponding random effect (better terminology: **error**) which acts as the experimental error on that level.

General Situation

- This means:
 - Start on the whole-plot level and forget about the split-plots.
 - Write down the corresponding model equation (incl. random effect / error).
 - Move on to the next level, expand equation with new terms (the upper level is now a block)
 - Etc.
- In R we just have to make sure that we tell `lmer` the correct random effects.
- In R it is sometimes useful to define new variables which identify the different experimental units on the different levels.

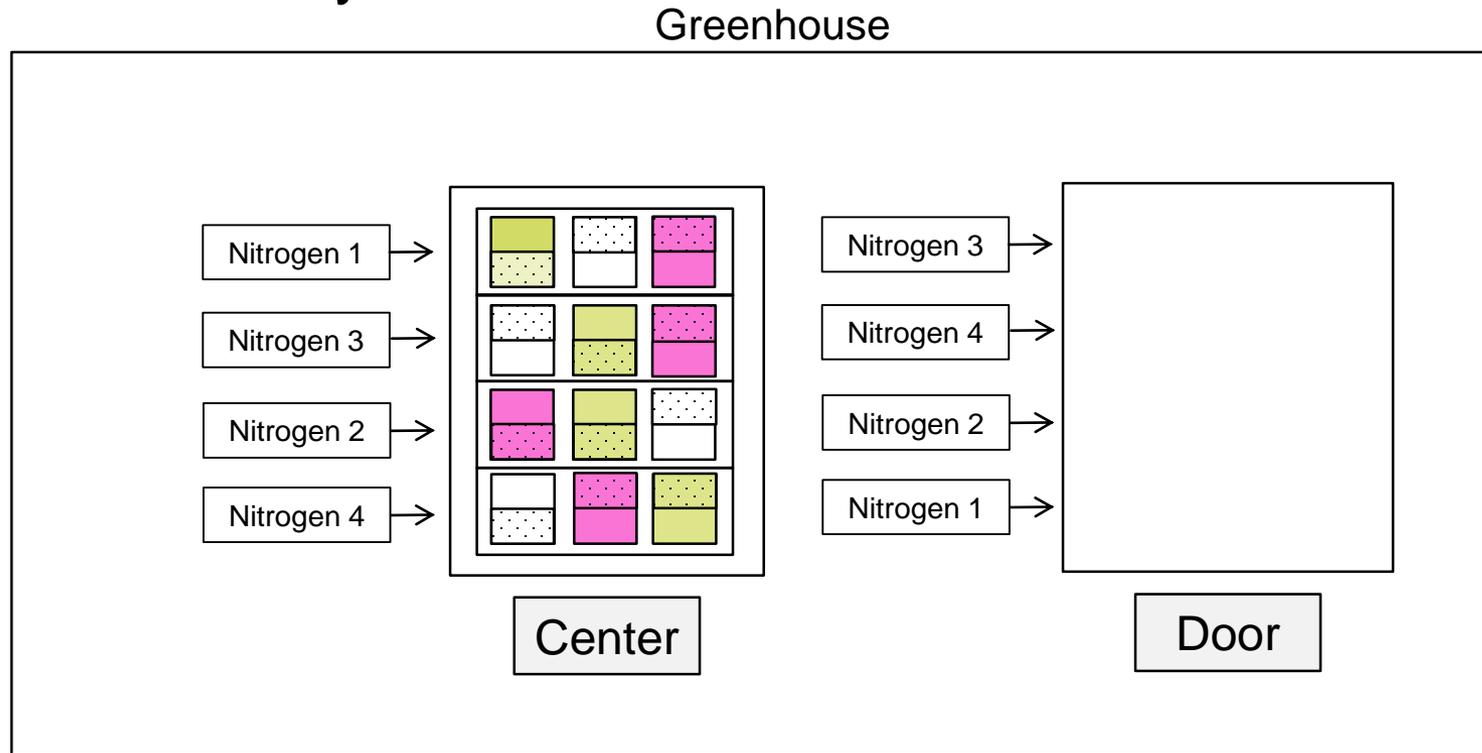
Example IV: Weed Biomass in Wetlands (Oehlert, 2000, Ex. 16.7)



- Experiment studies the effect of
 - **nitrogen** (4 levels of nitrogen)
 - **weed** (3 levels)
 - **clipping treatments** (2 levels: clipping / no clipping)on **plant growth** in wetlands.
- Experiment was performed as follows:
 - **8 trays**, whereof each holds three artificial wetlands (rectangular wire baskets)
 - 4 of the trays were placed on **a table near the door** of the greenhouse
 - 4 of the trays on **a table in the center** of the greenhouse
 - On **each table**, we randomly assign one of the trays to each of the **4 nitrogen treatments**.
 - Within **each tray**, we randomly assign the **3 weed treatments**.
 - In addition, each wetland is split in half. One half is chosen at random and will be clipped, the other half is not clipped.
 - After 8 weeks: measure fraction of biomass that is nonweed.

Example IV: Weed Biomass in Wetlands

Experimental layout



Example IV: Weed Biomass in Wetlands

- Let us follow the path of randomization:
 - Position in the greenhouse is a **block factor** (center / door)
 - **Trays** are **whole plots**, and **nitrogen level** is the **whole-plot factor**.
 - **Wetlands** are **split plots** and **weed treatment** is the **split-plot factor**.
 - **Wetland halves** are so called **split-split plots** and **clipping** is the **split-split-plot factor**.
- Hence, we have a so-called **split-split plot**.
- Let us now try to fit a model to this data-set in R.

Example IV: Weed Biomass in Wetlands

- We use the following model

```
> fit <- lmer(pct.nonweed.biomass ~ table + nitrogen + (1 | tray) +  
+           weed * nitrogen + (1 | wetland) +  
+           weed * nitrogen * clipping, data = wetland)  
> anova(fit)
```

Analysis of Variance Table of type III with Satterthwaite approximation for degrees of freedom

| | Sum Sq | Mean Sq | NumDF | DenDF | F.value | Pr(>F) | |
|------------------------|---------|---------|-------|---------|---------|-----------|-----|
| table | 0.16 | 0.16 | 1 | 3.0001 | 0.15 | 0.72113 | |
| nitrogen | 36.73 | 12.24 | 3 | 3.0001 | 11.46 | 0.03765 | * |
| weed | 1186.82 | 593.41 | 2 | 8.0000 | 555.45 | 2.613e-09 | *** |
| clipping | 125.45 | 125.45 | 1 | 12.0001 | 117.43 | 1.493e-07 | *** |
| nitrogen:weed | 157.57 | 26.26 | 6 | 8.0000 | 24.58 | 9.664e-05 | *** |
| weed:clipping | 0.25 | 0.12 | 2 | 12.0001 | 0.11 | 0.89246 | |
| nitrogen:clipping | 0.73 | 0.24 | 3 | 12.0001 | 0.23 | 0.87419 | |
| nitrogen:weed:clipping | 4.82 | 0.80 | 6 | 12.0001 | 0.75 | 0.62033 | |

- All **main-effects** and the **nitrogen × weed interaction** are significant.

Example IV: Weed Biomass in Wetlands

- We are here performing 3 experiments in 1.
- On the **whole-plot level** we have the “experiment”

| <i>Source</i> | <i>df</i> |
|-------------------------|--------------------|
| Table (block) | 1 |
| Nitrogen | 3 |
| Error (per tray) | 3 (= 7 - 4) |
| Total | 7 (= 8 - 1) |

- On the **split-plot level** we have the “experiment”

| <i>Source</i> | <i>df</i> |
|----------------------------|----------------------|
| Block (=Tray) | 7 |
| Weed | 2 |
| Weed × Nitrogen | 6 |
| Error (per wetland) | 8 (= 23 - 15) |
| Total | 23 (= 24 - 1) |

Example IV: Weed Biomass in Wetlands

- On the **split-split-plot level** we have the “experiment”

| <i>Source</i> | <i>df</i> |
|---------------------------------|-----------------------|
| Block (= wetland) | 23 |
| Clipping | 1 |
| Weed × Clipping | 2 |
| Nitrogen × Clipping | 3 |
| Nitrogen × Weed × Clipping | 6 |
| Error (per wetland half) | 12 (= 47 – 35) |
| Total | 47 (= 48 – 1) |

Summary

- Split plot designs and more complicated versions thereof are useful if some factors are harder (more expensive, ...) to vary than others.
- To identify the correct design we have to know the randomization procedure.
- The general situation can be very complex, but by following the different randomization levels/steps, setting up a model is easy.
- Mixed effects software like `lmer` automatically identifies the correct denominator for tests if the random effects / errors are stated correctly.