



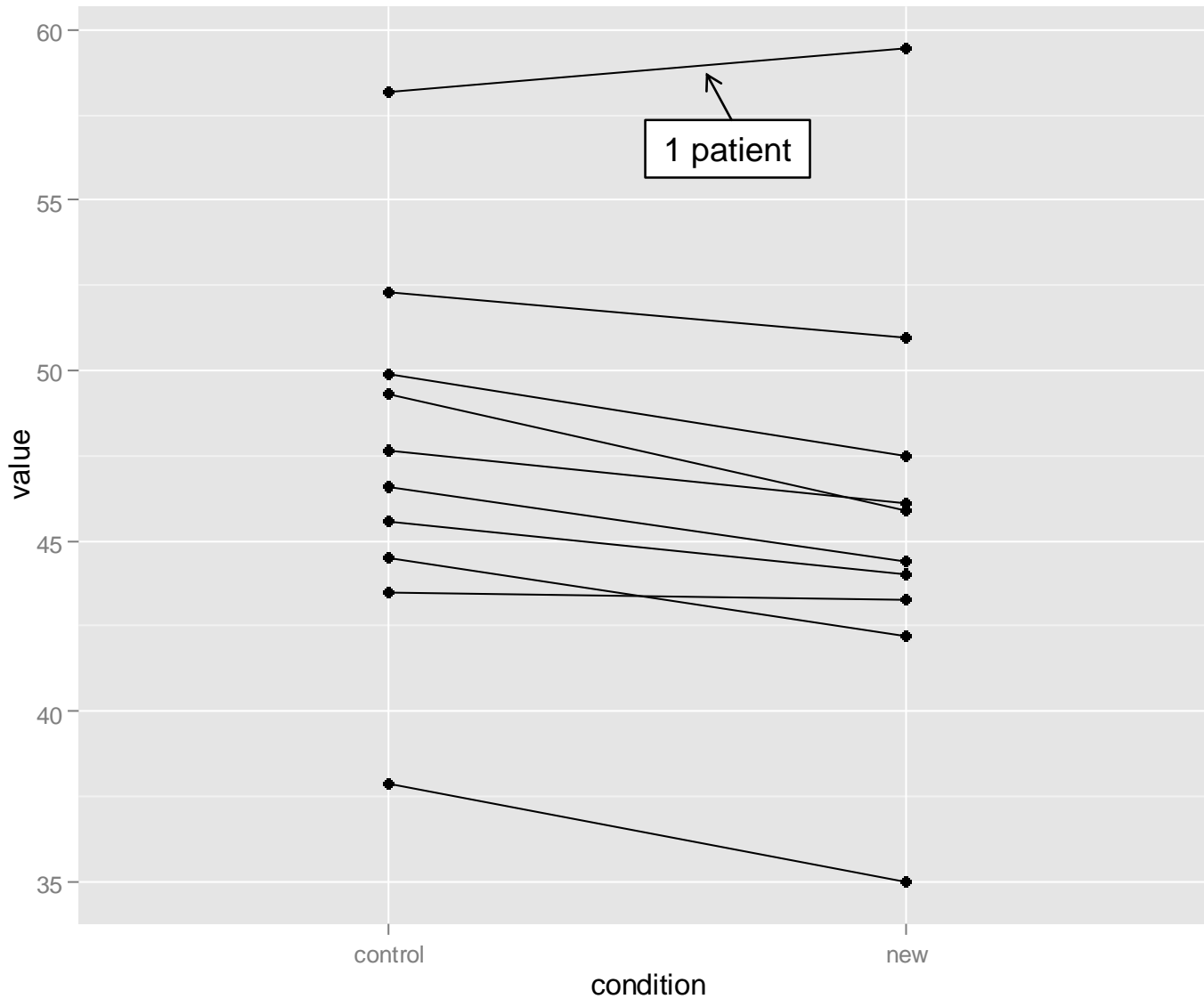
Complete Block Designs

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Remember: Paired t -Test (Example from Elliott, 2006)

- Want to compare two different eye-drops (“new” vs. “control”).
- **Every** subject gets **both treatments** (meaning: one per eye; at the **same** time).
- At the end, measure redness on quantitative scale in **every** eye.
- For every patient, calculate the difference “new - control”.
- Perform standard one-sample t -test with these differences.

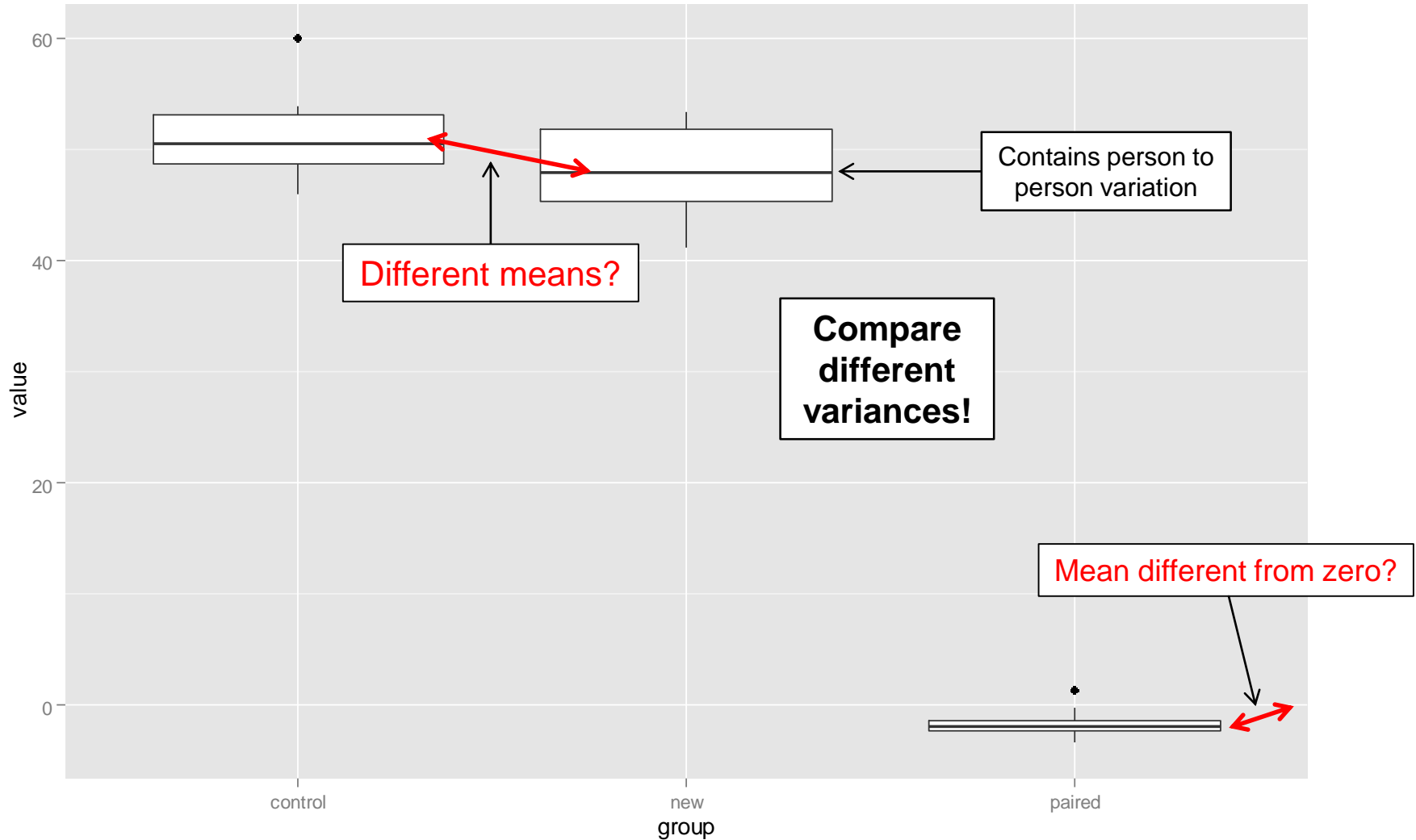
Fictional Data Set of 10 Patients



Paired t -Test

- Instead of using both eyes of 10 patients we could also do a similar experiment with
 - 10 patients getting the control treatment in one eye
 - 10 other patients getting the new treatment in one eye
- See next slide for potential data-sets.
- As mentioned in the first week, we can reduce **variance** by using **homogeneous experimental units**.
- A set of units that is homogeneous in some sense is called a **block**.
- In this example, a **block** is given by a **person**.

Paired vs. Unpaired Data



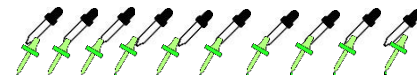
control treatment:
10 patients



new treatment:
10 (other) patients

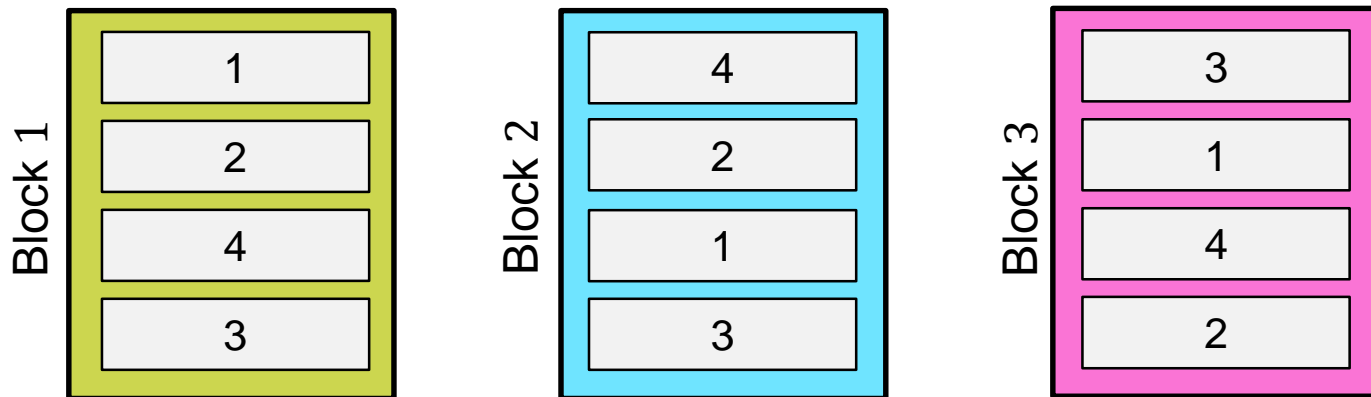


difference of treatments:
within 10 (other) patients



Randomized Complete Block Designs (RCB)

- A **Randomized Complete Block Design** (RCB) is the most basic blocking design.
- Assume we have r blocks containing g units each.



- Here, $r = 3$ blocks with $g = 4$ units.
- In every of the r blocks we randomly assign the g treatments to the g units, **independently** of the other blocks.

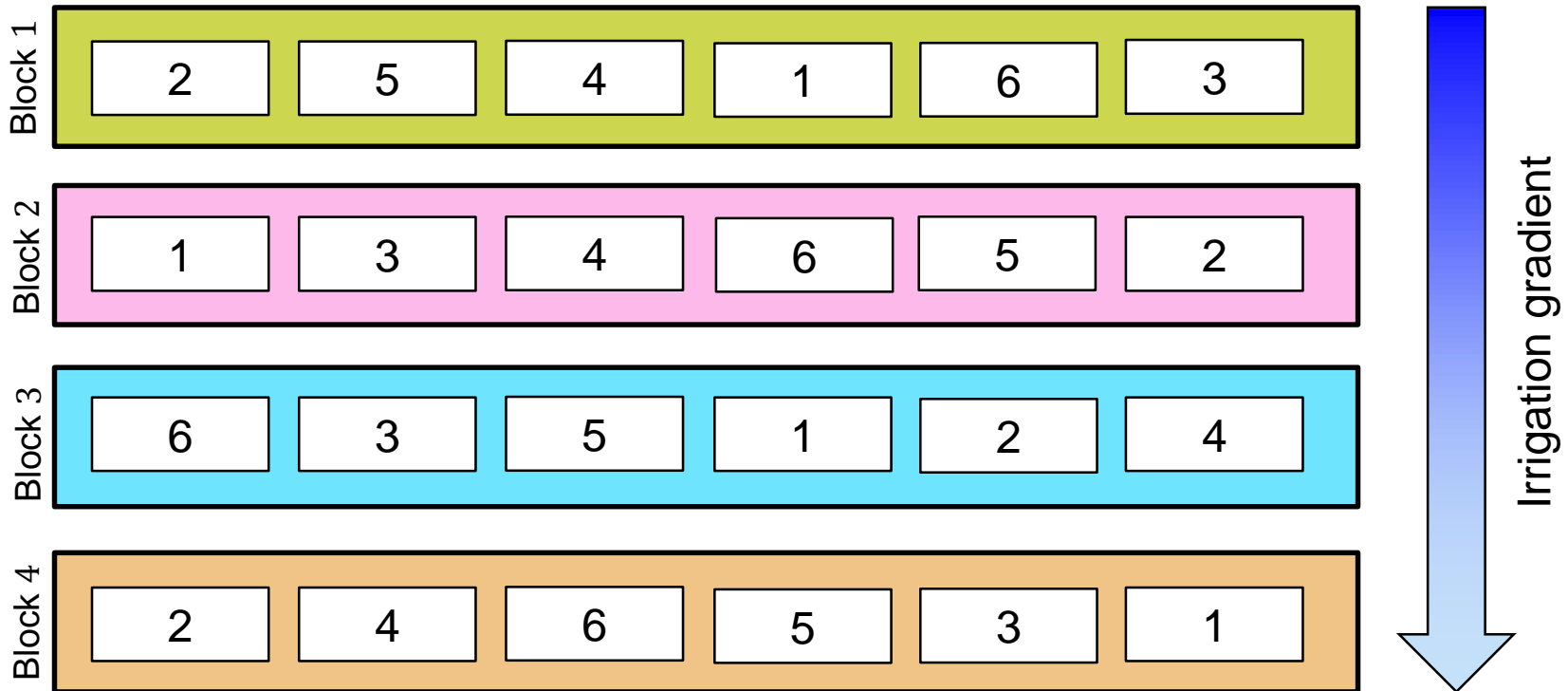
Randomized Complete Block Designs (RCB)

- Hence, a blocking design uses a **restricted randomization** scheme. Each block gets its “**own**” **randomization**.
- Blocking exists at the time of randomization!
- We call a blocking design **complete** if every treatment is used in every block.
- In the standard setup, we observe every treatment (only) **once** in every block, hence we have a total of r (the number of blocks) observations per treatment.
- Therefore, we have **no replicates** (for treatment and block combinations).

Example (Example 8.1 in Kuehl, 2000)

- Researchers wanted to evaluate the effect of several different fertilization timing schedules on stem tissue nitrate amounts.
- Treatment: Six different nitrogen application timing and rate schedules (including a control treatment of no nitrogen).
- Response: Stem tissue nitrate amount.
- Experiment design: irrigated field with a **water gradient** along one direction, see next slide.
- We already know:
Available moisture will have an influence on the response.

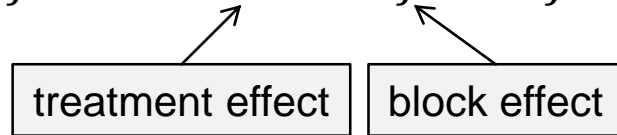
Example: Layout of Experimental Design



- Any differences in plant responses caused by the **water gradient** will be associated with **blocks**.
- We also say: we **control for the water gradient**.

Example: Analysis

- $Y_{ij} = \mu + \alpha_i + \beta_j + \epsilon_{ij}$ with the usual assumptions for ϵ_{ij} .



- By only using main effects we implicitly assume that the effects are **additive**.
- Due to the **balanced design** we can use our standard estimates (one at a time) and sum of squares.

```
> fit <- aov(y ~ block + treatment, data = nitro)
> summary(fit)
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)	
block	3	197.0	65.67	9.120	0.00112	**
treatment	5	201.3	40.26	5.592	0.00419	**
Residuals	15	108.0	7.20			

- Typically, we are **not** making inference about blocks (we already know that blocks are different!).

Interaction of Treatment with Block Factor

- The blocking may result in (very) large differences between units from different blocks (which is **ok**).
- In the model we **assumed** that the effects are additive.
- Meaning: the treatment effects are **constant** from block to block.
- If we only have **one observation** per treatment and block combination we can potentially only detect interaction effects of the multiplicative form.
- If we want to fit a model with interaction, we would need more than one observation per treatment and block combination. What does interaction mean?

Factorials in Complete Block Designs

- Conceptually it is straightforward to have (e.g.) a **two-factor factorial** in a randomized complete block design.
- The analysis is straightforward. In R we would just use the model formula `Y ~ Block + A * B`

<i>Source</i>	<i>df</i>	
Block	$r - 1$	
<i>A</i>	$a - 1$	
<i>B</i>	$b - 1$	
<i>AB</i>	$(a - 1) \cdot (b - 1)$	
Error	$(ab - 1) \cdot (r - 1)$	← “Leftovers”
Total	$rab - 1$	← # observations - 1

- We can test the interaction *AB* even if we only have one replicate per *AB* combination per block.

How Much Does Blocking Increase Precision?

- Squared standard errors for treatment means are
 - RCB design (what we've just done): $\frac{\sigma_{RCB}^2}{r}$
 - Completely randomized design: $\frac{\sigma_{CRD}^2}{n}$
- If we want to have the **same precision**, we have to ensure that

$$\frac{\sigma_{RCB}^2}{r} = \frac{\sigma_{CRD}^2}{n}.$$

If we know σ_{RCB}^2 and σ_{CRD}^2 than we have to use a ratio of

$$\frac{n}{r} = \frac{\sigma_{CRD}^2}{\sigma_{RCB}^2}.$$

Number of observations per treatment

How Much Does Blocking Increase Precision?

- σ_{RCB}^2 is estimated by MS_E of our RCB.
- What about σ_{CRD}^2 ?
- Can be estimated using a properly **weighted average** of MS_E and MS_{Block}

$$\hat{\sigma}_{CRD}^2 = w \cdot MS_{Block} + (1 - w) \cdot MS_E$$

where w is some weight (see Oehlert, page 323).

- **Relative efficiency** is then defined as:

$$RE = \frac{\hat{\sigma}_{CRD}^2}{\hat{\sigma}_{RCB}^2}$$

(sometimes multiply with correction factor for df's).

- RE gives us the ratio $\frac{n}{r}$.

How Much Does Blocking Increase Precision?

- In our example: relative efficiency ≈ 2 .
- Meaning: A CRD would need twice as many experimental units to achieve the **same efficiency** (precision).
- Here: 8 replications per treatment (instead of 4).
- **Easier for a quick check:** Have a look at the ratio $\frac{MS_{Block}}{MS_E}$






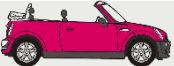


$$\frac{MS_{Block}}{MS_E} > 1 \Leftrightarrow \text{Relative Efficiency} > 1$$

More than One Blocking Factor

- Up to now: one blocking factor involved, i.e. we can block on a **single source** of variation.
- Sometimes: need to block on **more than one** source.
- We will discuss some special cases.
 - Latin Squares
 - Graeco-Latin Squares

Example: Car Tires (Kuehl, 2000, Example 8.2)

- An experiment tests **4 car tire treatments** (A, B, C, D) on **4 cars**. Response: Wear of a tire.
- Each treatment appears on one of the **4 positions** of each **car**. ← Block factors →
- Experiment set-up was as follows:

Tire position				
	<i>A</i>	<i>B</i>	<i>C</i>	<i>D</i>
	<i>B</i>	<i>C</i>	<i>D</i>	<i>A</i>
	<i>C</i>	<i>D</i>	<i>A</i>	<i>B</i>
	<i>D</i>	<i>A</i>	<i>B</i>	<i>C</i>

Latin Squares

- This design is a so called **Latin Square**.
- Each treatment (the Latin letters) appears exactly **once** in **each row** and exactly **once** in **each column**.
- A Latin Square **blocks** on both **rows** and **columns simultaneously**.
- The design is very **restrictive**. A Latin Square needs to have
 - g treatments (the Latin letters)
 - Two block factors each having g levels (the rows and the columns)
 - Hence, a total of g^2 experimental units
- We're only seeing g^2 out of g^3 possible combinations (but the subset we see is selected in a smart, balanced way).

Latin Squares

- A Latin Square is nothing else than an assignment of treatments to units with the **side constraints**
 - each treatment appears **exactly once in each row**.
 - each treatment appears **exactly once in each column**.
- Picking a random Latin Square isn't trivial: Fisher-Yates algorithm (see book for details).

Analysis of Latin Squares

- Use **main effects model** with **treatment, row and column effects**.

$$Y_{ijk} = \mu + \alpha_i + \beta_j + \gamma_k + \epsilon_{ijk}$$

Diagram illustrating the components of the model equation:






- α_i is associated with **treatment**.
- β_j is associated with **Block factor 1 (rows)**.
- γ_k is associated with **Block factor 2 (columns)**.

- The design is **balanced** having the effect that our usual estimators and sums of squares are “working”.
- As in an RCB we do **not** test for the block effects.
- Latin Squares can have **few degrees** of freedom for error if g is small, making detection of treatment effects difficult:

g	df of MS_E
3	2
4	6
5	12





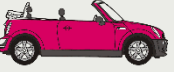


Latin Squares

- Just because the design contains the word “square” doesn’t mean that the physical layout of the experiment has to be a square.
- Often, one blocking factor is **time**: Think of testing 5 different machines (*A, B, C, D, E*) on 5 days with 5 operators (response: yield of machine):

<i>Operator</i>					
<i>Mon</i>	<i>E</i>	<i>B</i>	<i>C</i>	<i>A</i>	<i>D</i>
<i>Tue</i>	<i>B</i>	<i>D</i>	<i>E</i>	<i>C</i>	<i>A</i>
<i>Wed</i>	<i>A</i>	<i>C</i>	<i>D</i>	<i>B</i>	<i>E</i>
<i>Thu</i>	<i>C</i>	<i>E</i>	<i>A</i>	<i>D</i>	<i>B</i>
<i>Fri</i>	<i>D</i>	<i>A</i>	<i>B</i>	<i>E</i>	<i>C</i>

Graeco Latin Squares

- What if we have one more blocking criterion?
- Use so called **Graeco Latin Squares** (if applicable).
- Take a Latin Square and superimpose it with another block factor, denoted by Greek letters (here: think of driver)

<i>Car</i>				
	$A\alpha$	$B\gamma$	$C\delta$	$D\beta$
	$B\beta$	$A\delta$	$D\gamma$	$C\alpha$
	$C\gamma$	$D\alpha$	$A\beta$	$B\delta$
	$D\delta$	$C\beta$	$B\alpha$	$A\gamma$

Graeco Latin Squares

- The Latin letters occur once in each row and column
- The Greek letters occur once in each row and column
- In addition: each Latin letter occurs exactly **once** with each Greek letter.
- Use **main effects model**

Latin squares

$$Y_{ijkl} = \mu + \alpha_i + \beta_j + \gamma_k + \delta_l + \epsilon_{ijkl}$$

to analyze data.

treatment

Block factor 1
(rows)

Block factor 2
(columns)

Block factor 3
(Greek letters)

More General Situations

- In practice, (Graeco) Latin Squares are often **impractical** due to the **very restrictive assumptions** on the **number of levels** of the involved treatment and block factors.
- E.g., think of the car tire example with 7 instead of 4 tire treatments.
- Or going back to the intro example: What if we wanted to compare **three** different eye-drops?
- This will lead us to **balanced incomplete block designs (BIBD)**, see later.



General Rules for Analyzing Block Designs

- As we have seen, we treat block factors just as other factors in our model formulas.
- Typically, a block effect is assumed to be **additive** (i.e., main effects only).
- Block factors are **not** tested but they can be examined with respect to efficiency gain.
- ANOVA table and df's are “as usual”.
- A possible interaction between block and treatment factor(s) is difficult to sell.