## Power, Type I and II Error

- Type I error $=$ reject $H_{0}$ when $H_{0}$ is true. The probability of a Type I error is called the significance level of the test, denoted by $\alpha$.
- Type II error= fail to reject $\mathrm{H}_{0}$ when $\mathrm{H}_{0}$ is false. The probability of a type II error is denoted by $\beta$.
- The power of a test is

$$
\text { power }=P\left(\text { reject } \mathrm{H}_{0} \mid \mathrm{H}_{0} \text { is false }\right)=1-\beta
$$

Test statistic under $H_{0}$ and $H_{A}$


$$
\left(t^{*}=t_{1-\alpha / 2}\right)
$$

The power depends on $\alpha, \delta, \sigma$ and $n$

## Power calculation in general

- Prospective: want a power of $\geq 80 \%$, determine the necessary sample size.
- Retrospective: sample size was given, test not significant, how much power did we have?


## 2-sample $t$ test

Let $X_{11}, \ldots, X_{1 n}$ iid and $X_{21}, \ldots, X_{2 n}$ iid independent.
$H_{0}: X_{1 i} \sim \mathcal{N}\left(\mu_{1}, \sigma^{2}\right), X_{2 j} \sim \mathcal{N}\left(\mu_{2}, \sigma^{2}\right)$ with $\mu_{1}=\mu_{2}$
$H_{A}: X_{1 i} \sim \mathcal{N}\left(\mu_{1}, \sigma^{2}\right), X_{2 j} \sim \mathcal{N}\left(\mu_{2}, \sigma^{2}\right)$ with $\mu_{1} \neq \mu_{2}$
Under $H_{0}$ :
$\bar{X}_{1}-\bar{X}_{2} \sim \mathcal{N}\left(0, \sigma^{2}\left(\frac{1}{n}+\frac{1}{n}\right)\right) \Rightarrow \frac{\bar{X}_{1}-\bar{X}_{2}}{\sigma \sqrt{2 / n}} \sim \mathcal{N}(0,1)$
Estimate $\sigma^{2}$ by $S_{p}^{2}=\frac{S_{1}^{2}+S_{2}^{2}}{2}$
$t=\frac{\bar{X}_{1}-\bar{X}_{2}}{S_{p} \sqrt{2 / n}}$ follows a $t$ distribution with $2 n-2 \mathrm{df}$

## Power calculation

We reject $\mathrm{H}_{0}$ if $t=\frac{\left|\bar{x}_{1}-\bar{x}_{2}\right|}{s_{p} \sqrt{2 / n}}>t_{1-\alpha / 2,2 n-2}$.

$$
1-\beta=P\left(\left.\frac{\bar{x}_{1}-\bar{X}_{2}}{S_{P} \sqrt{2 / n}}<-t_{1-\alpha / 2,2 n-2} \right\rvert\, H_{A}\right)+P\left(\left.\frac{\bar{x}_{1}-\bar{x}_{2}}{S_{P} \sqrt{2 / n}}>t_{1-\alpha / 2,2 n-2} \right\rvert\, H_{A}\right) .
$$

Under $\mathrm{H}_{A} \frac{\bar{x}_{1}-\bar{X}_{2}-\delta}{S_{p} \sqrt{2 / n}}$ follows a $t$ distribution with $2 n-2 \mathrm{df}$.
This implies

$$
1-\beta=P\left(\frac{\bar{X}_{1}-\bar{X}_{2}-\delta}{S_{p} \sqrt{2 / n}}>t_{1-\alpha / 2}-\frac{\delta}{S_{p} \sqrt{2 / n}}\right)+\underbrace{P\left(\frac{\bar{X}_{1}-\bar{X}_{2}-\delta}{S_{p} \sqrt{2 / n}}<t_{\alpha / 2}-\frac{\delta}{S_{p} \sqrt{2 / n}}\right.}_{\text {Prob } \approx 0}) .
$$

## Quantiles of the t distribution



It follows that $t_{\beta}=t_{1-\alpha / 2}-\frac{\delta \sqrt{n}}{S_{p} \sqrt{2}}$

## Equations for power calculation

For any $\delta \neq 0$, the following equations hold.

$$
\begin{align*}
t_{\beta} & =t_{1-\alpha / 2}-\frac{|\delta| \sqrt{n}}{s_{p} \sqrt{2}}  \tag{1}\\
n & =2\left(t_{1-\alpha / 2}-t_{\beta}\right)^{2} \cdot \frac{s_{p}^{2}}{\delta^{2}} \tag{2}
\end{align*}
$$

## One-way anova

- The power of the F test for $\mathrm{H}_{0}: \mu_{1}=\mu_{2}=\ldots=\mu_{I}$ is

$$
1-\beta=P_{H_{A}}(\text { Test significant })=P\left(F>F_{1-\alpha, I-1, N-I} \mid H_{A}\right) .
$$

- The distribution of $F$ under $H_{A}$ follows a noncentral $F$ distribution with non-centrality parameter $\delta^{2}=\frac{J \sum_{\sigma^{2}} A_{i}^{2}}{}$ and $I-1$ and $N-I$ degrees of freedom.
- There are tables, graphs and software (e.g. GPower) which determine the power given $I-1, N-I, \alpha$ and $\delta$.
- Use $\Delta=\frac{\max A_{i}-\min A_{i}}{\sigma}$.

Detectable differences $\Delta$ for $\alpha=5 \%$ and $1-\beta=90 \%$

|  | Number of groups $I$ |  |  |  |  |
| ---: | :---: | :---: | :---: | :---: | :---: |
| $J$ | 2 | 3 | 4 | 5 | 6 |
| 2 | 6.796 | 6.548 | 6.395 | 6.333 | 6.317 |
| 3 | 3.589 | 3.838 | 3.967 | 4.065 | 4.149 |
| 4 | 2.767 | 3.010 | 3.148 | 3.251 | 3.337 |
| 5 | 2.348 | 2.568 | 2.698 | 2.795 | 2.876 |
| 6 | 2.081 | 2.280 | 2.401 | 2.492 | 2.567 |
| 7 | 1.890 | 2.073 | 2.186 | 2.271 | 2.341 |
| 8 | 1.745 | 1.915 | 2.020 | 2.100 | 2.166 |
|  |  |  |  |  |  |
| 10 | 1.534 | 1.684 | 1.778 | 1.850 | 1.910 |
| 12 | 1.385 | 1.521 | 1.607 | 1.673 | 1.727 |
| 14 | 1.273 | 1.398 | 1.478 | 1.539 | 1.589 |
| 16 | 1.185 | 1.301 | 1.375 | 1.432 | 1.479 |
| 18 | 1.112 | 1.222 | 1.292 | 1.345 | 1.390 |
| 20 | 1.052 | 1.155 | 1.222 | 1.273 | 1.315 |
| 22 | 1.000 | 1.099 | 1.162 | 1.210 | 1.251 |
| 24 | 0.956 | 1.050 | 1.110 | 1.157 | 1.195 |
| 26 | 0.917 | 1.007 | 1.065 | 1.109 | 1.146 |
| 28 | 0.882 | 0.969 | 1.025 | 1.068 | 1.103 |
| 30 | 0.851 | 0.935 | 0.989 | 1.030 | 1.065 |
| 40 | 0.734 | 0.806 | 0.852 | 0.888 | 0.918 |
| 60 | 0.597 | 0.655 | 0.693 | 0.722 | 0.747 |
| 80 | 0.516 | 0.566 | 0.599 | 0.624 | 0.645 |
| 100 | 0.461 | 0.506 | 0.535 | 0.558 | 0.577 |
| 200 | 0.325 | 0.357 | 0.377 | 0.393 | 0.407 |
| 500 | 0.205 | 0.225 | 0.238 | 0.248 | 0.257 |
| 1000 | 0.145 | 0.159 | 0.168 | 0.176 | 0.181 |

## Daily weight gains

Average daily weight gains are to be compared among pigs receiving 4 levels of vitamin $\mathrm{B}_{12}$ in their diet.

We estimate $\sigma$ with $\hat{\sigma}=0.015 \mathrm{lbs} . /$ day and we would like to detect a difference $\max A_{i}-\min A_{i}=0.03 \mathrm{lbs} /$ day. We set $\alpha=0.05$ and want a power of 0.90 at least for a balanced design. This implies
$\Delta=2$ and leads to a minimum of $n=9$ pigs per group.

