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Collinearity = Correlated Predictors

If ≥ 2 predictors are strongly correlated, i.e. explain very similar aspects of the response, OLS estimation is difficult. The regression coefficients will be less precise, and the interpretation of the results is more difficult.

There is a need to recognize collinearity!

1) Plot the correlation matrix of the predictors plotcorr(cor(my.dat))

2) Variance Inflation Factors

$$Var(\hat{\beta}_{k}) = \sigma_{E}^{2} \cdot \frac{1}{1 - R_{k}^{2}} \cdot \frac{1}{\sum_{i=1}^{n} (x_{ik} - \overline{x}_{k})}$$

How to Deal with Correlated Predictors?

- 1) Generate new variables
 - \rightarrow see example on next slides...
- 2) Variable selection

Only use the relevant predictors, and omit the redundant ones. This often helps a lot. We will be discussing variable selection in detail.

3) The Lasso and Ridge Regression

These are penalized OLS regression methods, which sparsely spend degrees of freedom. To be discussed later.

Example

Understanding how car drivers adjust their seat would greatly help engineers to design better cars. Thus, the measured

hipcenter = horizontal distance of hips to steering wheel

and tried to explain it with several predictors, namely:

Age	age in years
Weight	weight in pounds
HtShoes, Ht, Seated	height w/o, w/ shoes, seated height
Arm, Thigh, Leg	arm, thigh and leg length

We first fit a model with all these (correlated!) predictors

Example: Fit with All Predictors

- > library(faraway); data(seatpos)
- > summary(lm(hipcenter~., data=seatpos))

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	436.43213	166.57162	2.620	0.0138	*
Age	0.77572	0.57033	1.360	0.1843	
Weight	0.02631	0.33097	0.080	0.9372	
HtShoes	-2.69241	9.75304	-0.276	0.7845	
Ht	0.60134	10.12987	0.059	0.9531	
Seated	0.53375	3.76189	0.142	0.8882	
Arm	-1.32807	3.90020	-0.341	0.7359	
Thigh	-1.14312	2.66002	-0.430	0.6706	
Leg	-6.43905	4.71386	-1.366	0.1824	

Residual standard error: 37.72 on 29 degrees of freedom Multiple R-squared: 0.6866, Adjusted R-squared: 0.6001 F-statistic: 7.94 on 8 and 29 DF, p-value: 1.306e-05

Collinearity in the Seat Position Example

> vif(fit)

Age	Weight	HtShoes	Ht
1.997931	3.647030	307.429378	333.137832
Seated	Arm	Thigh	Leg
8.951054	4.496368	2.762886	6.694291



VIF \geq 5 is critical, VIF \geq 10 is dangerous. The observed values mean that the standard errors of the estimates are inflated by a factor of about 18x.

Example: Generating New Variables

The body height is certainly a key predictors when it comes to the position of the driver seat. We leave this as it was, and change several of the other predictors:

age <- Age
bmi <- (Weight*0.454)/(Ht/100)^2
shoes <- HtShoes-Ht
seated <- Seated/Ht
arm <- Seated/Ht
thigh <- Thigh/Ht
leg <- Leg/Ht</pre>

Does this solve the correlation problem...?

Example: New Correlation Matrix

> vif(fit00)

age bmi height shoes seated 1.994473 1.408055 1.968447 1.155285 1.851884 arm thigh leg 2.044727 1.284893 1.480397



Example: Fit with New Predictors

> summary(lm(hipc~., data=new.seatpos))

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	-632.0063	490.0451	-1.290	0.207	
age	-0.7402	0.5697	-1.299	0.204	
bmi	-0.4234	2.2622	-0.187	0.853	
height	3.6521	0.7785	4.691	5.98e-05	* * *
shoes	2.6964	9.8030	0.275	0.785	
seated	-171.9495	631.3719	-0.272	0.787	
arm	180.7123	655.9536	0.275	0.785	
thigh	141.2007	443.8337	0.318	0.753	
leg	1090.0111	806.1577	1.352	0.187	

Residual standard error: 37.71 on 29 degrees of freedom Multiple R-squared: 0.6867, Adjusted R-squared: 0.6002 F-statistic: 7.944 on 8 and 29 DF, p-value: 1.3e-05

Variable Selection: Why?

We want to fit a regression model...

- **Case 1:** functional form and predictors exactly known → estimation, test, confidence and prediction intervals
- Case 2: neither functional form nor the predictors are known → explorative model search among potential predictors
- **Case 3:** we are interested in only 1 predictor, but want to correct for the effect of other covariates

 \rightarrow which covariates we need to correct for?

Question in cases 2 & 3: WHICH PREDICTORS TO USE?

Variable Selection: Technical Aspects

We want to keep a model small, because of

1) Simplicity

 \rightarrow among several explanations, the simplest is the best

2) Noise Reduction

 \rightarrow unnecessary predictors leads to less accuracy

3) Collinearity

 \rightarrow removing excess predictors facilitates interpretation

4) Prediction

 \rightarrow less variables, less effort for data collection

Method or Process?

- Variable selection is not a method! The search for the best predictor set is an iterative process. It also involves *estimation*, *inference* and *model diagnostics*.
- For example, outliers and influential data points will not only change a particular model – they can even have an impact on the model we select. Also variable transformations will have an impact on the model that is selected.
- Some iteration and experimentation is often necessary for variable selection. *The ultimate aim is finding a model that is smaller, but as good or even better than the original one.*

Example: Mortality Data

```
> summary(fit.orig)
Coefficients:
```

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	1496.4915	572.7205	2.613	0.01224	*
JanTemp	-2.4479	0.8808	-2.779	0.00798	* *
•••					
Dens	11.9490	16.1836	0.738	0.46423	
NonWhite	326.6757	62.9092	5.193	5.09e-06	* * *
WhiteCollar	-146.3477	112.5510	-1.300	0.20028	

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Residual standard error: 34.23 on 44 degrees of freedom Multiple R-squared: 0.7719, Adjusted R-squared: 0.6994 F-statistic: 10.64 on 14 and 44 DF, p-value: 6.508e-10

Note: due to space constraints, this is only part of the output.

Backward Elimination with p-Values

- **Aim:** Reducing the regression model such that the remaining predictors show a significant relation to the response.
- **How:** We start with the full model and then exclude the least significant predictor in a step-by-step manner, as long as its p-value is greater than $\alpha_{crit} = 0.05$.
- In R:> fit <- update(fit, . ~ . RelHum)
 > summary(fit)
- \rightarrow Re-fit the model after each exclusion!
- → Wording: Backward Elimination with p-Values
- → For prediction, one also uses $\alpha_{crit} = 0.10/0.15/0.20$

Applied Statistical Regression AS 2012 – Week 09 Example: Final Result

> ft09 <- update(ft08, .~.-WhiteCollar); summary(ft09)</pre>

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	992.2069	79.6994	12.449	< 2e-16	* * *
JanTemp	-2.1304	0.5017	-4.246	8.80e-05	* * *
Rain	1.8122	0.5066	3.577	0.000752	* * *
Educ	-16.4207	6.1202	-2.683	0.009710	* *
NonWhite	268.2564	38.8832	6.899	6.56e-09	* * *
NOx	18.3230	4.3960	4.168	0.000114	* * *

Residual standard error: 33.47 on 53 degrees of freedom Multiple R-squared: 0.7373, Adjusted R-squared: 0.7125 F-statistic: 29.75 on 5 and 53 DF, p-value: 2.931e-14

 \rightarrow 9 predictors are eliminated, 5 remain in the final model.

Interpretation of the Result

- The remaining predictors are now "more significant" than before. This is almost always the case. Do not overestimate the importance of these predictors!
- Collinearity among the predictors is usually at the root of this observation. The predictive power is first spread out among several predictors, then it becomes concentrated.
- **Important**: the removed variables can still be related to the response. If we run a simple linear regression, they can even be significant. In the multiple linear model however, there are other, better, more informative predictors.

Alternatives to Backward Elimination

Backward elimination that is based on p-values requires laborious handwork (*in R*) and has a few disadvantages...

- When the principal goal is prediction, then the resulting models are often too small, i.e. there are bigger models which yield a more accurate prognosis.
- From a (theoretical) mathematical viewpoint variable selection via the AIC/BIC criterions is more suitable.
- In a step-by-step backward elimination, the best model is often missed. Evaluating more models can be very beneficial for finding *the best one*...

The AIC/BIC Criteria

Aim: Judging the quality of a regression model

→ Gauging Goodness-of-Fit vs. The Number of Predictors

AIC-Criterion:

$$AIC = -2 \max(\log likelihood) + 2p$$

$$= const + n \log(RSS / n) + 2p$$

BIC-Criterion:

$$BIC = -2\max(\log likelihood) + p\log n$$
$$= const + n\log(RSS / n) + p\log n$$

Backward Elimination with AIC/BIC

- **Aim:** Reducing the regression model such that the remaining predictors are *necessary* for describing the response.
- **How:** We start with the full model and then in a step-by-step manner exclude the predictor that leads to the biggest improvement in AIC/BIC.
- In R: > fit.aic <- step(fit, dir="backward", k=2)
 > fit.bic <- step(fit, dir="backward", k=log(59))</pre>
- → The variable selection stops when AIC/BIC cannot be improved anymore. There is neither a need nor a guarantee that the selected predictors are significant.

Example: Models with AIC/BIC

AIC:	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	1035.5384	85.1924	12.155	< 2e-16	* * *
JanTemp	-2.0188	0.5043	-4.003	0.000200	* * *
Rain	1.9637	0.5146	3.816	0.000363	* * *
Educ	-11.7708	6.9613	-1.691	0.096842	•
NonWhite	261.5379	38.8830	6.726	1.35e-08	* * *
WhiteCollar	-139.2913	102.0379	-1.365	0.178101	
NOx	19.4440	4.4372	4.382	5.73e-05	* * *
BIC:	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	992.2069	79.6994	12.449	< 2e-16	* * *
JanTemp	-2.1304	0.5017	-4.246	8.80e-05	* * *
Rain	1.8122	0.5066	3.577	0.000752	* * *
Educ	-16.4207	6.1202	-2.683	0.009710	* *
NonWhite	268.2564	38.8832	6.899	6.56e-09	* * *
NOx	18.3230	4.3960	4.168	0.000114	* * *

Applied Statistical Regression AS 2012 – Week 09 Visualization of Variable Selection

> plot(fit.aic\$anova\$AIC, ...)





AIC or BIC?

Usually, both criteria lead to similar models. BIC penalizes bigger models harder, with factor $\log n$ instead of factor 2.

→ "BIC models" tend to be smaller than "AIC models"!

Rule of the thumb for criterion choice:

- **BIC** is used when we are after a small model that is easy to interpret, i.e. in cases where understanding the predictor-response relation is the primary goal.
- **AIC** is used when the principal aim is the prediction of future observations. In these cases, small out-of-sample error is key, but neither the number or meaning of the predictors.

Alternative Search Heuristics

Forward Selection

- 1) Start with an empty model, i.e. only the intercept, but no predictors. The fitted value is the mean of the responses.
- 2) In a step-by-step manner, the predictor which leads to the best AIC/BIC value is added to the model.
- 3) Adding predictor variables is repeated until the AIC/BIC value can no longer be improved.
- **R:** > fit.aic <- step(fit, dir="forward", k=2)
- → Forward Selection is used with big datasets, where backward elimination is too time consuming.

Alternative Search Heuristics

Stepwise Model Search

- This is an alternation of forward and backward steps. We can either start with the full model (1. step is backwards) or with the empty model (1. step is forward).
- In each forward step, all predictors can be added, also these that were excluded before. In each backward step, any of the predictors can be kicked out of the model (again).
- → Similar to Backward Elimination resp. Forward Search
- \rightarrow Not much more time consuming, but more exhaustive
- → Recommended!

Stepwise Model Search in R

Starting with an empty model:

- > null <- lm(Mortality ~ 1, data=mortality)</pre>
- > all <- lm(Mortality ~ ., data=mortality)</pre>
- > fit <- step(null, scope=list(upper=all))</pre>

Starting with the full model:

> fit <- step(all, direction="both", k=2)</pre>

Note:

Argument scope=... allows specifying arbitrary minimal and maximal models for both cases. Then some predictors can be added or be removed from the model.

Alternative Search Heuristics

All Subsets Regression

- When m predictors are present, there are in fact 2^m different models that could be tried for finding the best one.
- In cases where *m* is small (i.e. $m \approx 10-20$) all submodels (up to a certain size) can be tried and evaluated by computing the AIC/BIC criterion.
- \rightarrow Complete search, but enormous computing time needed
- \rightarrow Yields a good solution, but not the causal model either
- \rightarrow Recommended for small dataset where it is feasible
- → R implementation with function leaps()

All Subsets Regression in R

R commands:

- > library(leaps)
- > summary(out)
- > plot(out)

Note:

The procedure starts with the empty model and for each number of predictors, identifies the nbest=1 models. By typing ~. in the formula, all predictors are allowed. The maximum model size that is search can be determined with nvmax=14.

Visualization of All Subsets Selection



BIC-Modellevaluation nach All Subsets Regression

Final Remarks

- Each search heuristics yields a different "best model".
- If we had another data sample from the same population and would repeat the variable selection using the same heuristic, the result might turn out to be different.
- The "best model" has the character of a random variable.

How to deal with that in practice?

We should not only consider the one "best model" according to a particular search heuristic, but a number of alternative model with similar performance (if they exist).

Model Selection with Hierarchical Input

 \rightarrow Some regression models have a natural hierarchy.

I.e. in polynomial models, x^2 is a higher order term than x

Important:

Lower order terms should not be removed from the model before higher order terms in the same variable. As an example, consider the polynomial model:

$$Y = \beta_0 + \beta_1 x + \beta_2 x^2 + \varepsilon$$

→ see blackboard...

Interactions and Categorical Input

Models with Interactions

Do not remove main effect terms if there are interactions with these predictors contained in the model.

Categorical Input

- If a single dummy coefficient is non-significant, we cannot just kick this term out of the model, but we have to test the entire block of indicator variables.
- When we work manually and testing based, this will be done with a partial F-test. When working criterion based, step() does the right thing

Cross Validation: Why?

- We have seen before that on a given dataset, a *bigger model* always yields a *better fit*, i.e. smaller residuals, and thus better RSS, R-squared, error variance, etc.
- Bigger models have an *unfair advantage* because there are *more predictors*. The AIC criterion tries to balance this by penalizing for the number of predictors used.
- If the *ultimate goal* is *predicting new datapoints*, then it is self suggesting to identify a model which does well at this. For making the right choice, we mimic the prediction task on our training sample with **Cross Validation**.

10-Fold Cross Validation

Idea:

- 0) Split the (training) data into 10 equally sized folds
- 1a) Use folds 1-9 for the fit, and use the model to predict fold 10
- 1b) On fold 10, the forecasting performance is measured by computing the RSS, i.e. the squared difference between the forecasted and true values
- 2) Use folds 1-8 & 10 for fitting, predict fold 9 and record RSS
- 3) Use folds 1-7 & 9-10 for fitting, predict fold 8 and record RSS

. . .

4)

10-Fold Cross Validation

Summary:

- Each observation is forecasted and gauged against the true values exactly 1x. On the other hand, it is used 9x during the model fitting process.
- With cross validation, we evaluate the "out-of-sample"-Performance, i.e. how precisely a model can forecast observations that were not used for fitting the model.
- In this regard, bigger and/or more complex models are not necessarily better than smaller/simpler ones.

Cross Validation

Further remarks:

- Cross validation is often used for identifying the most predictive model from a few candidate models that were found by stepwise variable selection procedures.
- There are alternatives to 10-fold CV. Popular is n-fold CV, which is known as Leave-One-Out Cross Validation.
- In R, it's easy to code "for-loops" that do the job, but there are also existing functions (that have some limits...):
 - > library(DAAG)
 - > CVlm(data, formula, fold.number, ...)