

# *Multi-Factor Experiments*

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- Twoway anova
- Interactions
- More than two factors

# *Hypertension: Effect of biofeedback*

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Biofeedback	Biofeedback	Medication	Control
+ Medication			

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158	188	186	185
163	183	191	190
173	198	196	195
178	178	181	200
168	193	176	180

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# Main effects

Treatment means:

	Medication		Total
	no	yes	
Biofeedback			
no	190	186	188
yes	188	168	178
Total	189	177	183

**main effect** of biofeedback:  $188 - 178 = 10$  mmHg

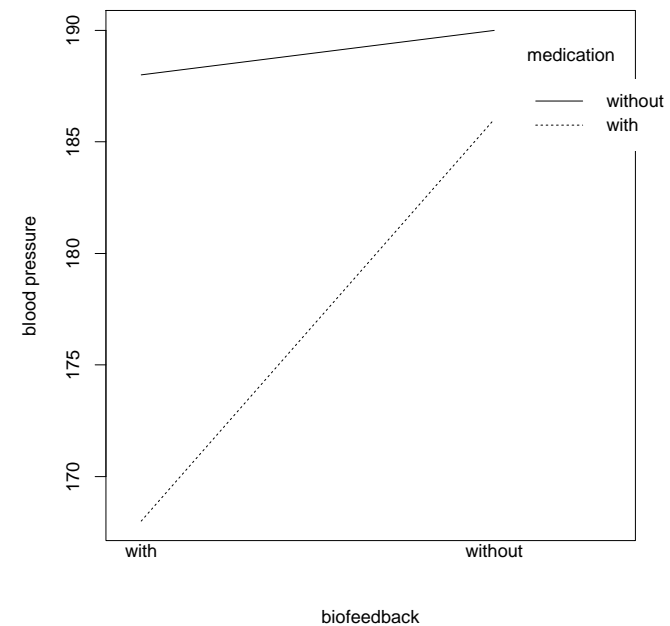
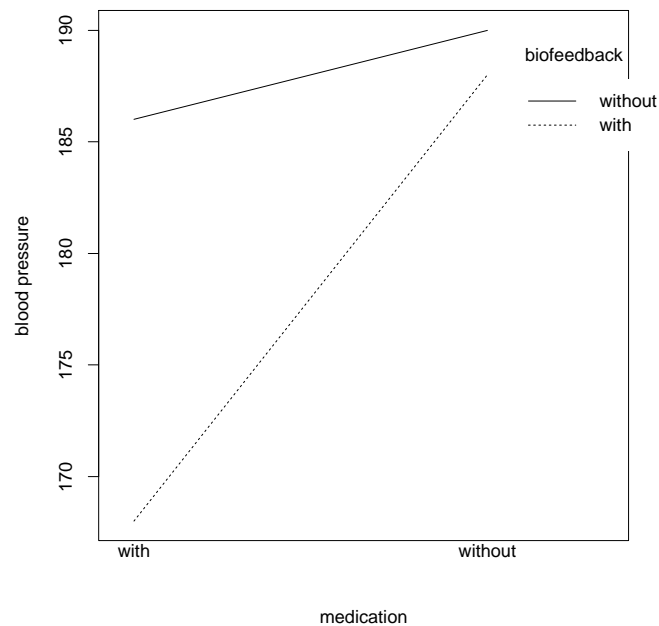
# Interaction plot

Effect of biofeedback with medi: 18 mmHg

≠

Effect of biofeedback without medi: 2 mmHg

→ Interaction



# Model for two factors

$$Y_{ijk} = \mu + A_i + B_j + (AB)_{ij} + \epsilon_{ijk}$$

$$i = 1, \dots, I; j = 1, \dots, J; k = 1, \dots, n.$$

$$\sum A_i = 0, \sum B_j = 0, \sum_i (AB)_{ij} = \sum_j (AB)_{ij} = 0.$$

$A_i$  :  $i$ th effect of factor A

$B_j$  :  $j$ th effect of factor B

$\mu + A_i + B_j$  : overall mean + effect of factor A on level  $i$   
+ effect of factor B on level  $j$

$(AB)_{ij}$  : deviation from additive model

# Parameter estimation

$$\hat{\mu} = y_{...}, \hat{A}_i = y_{i..} - y_{...} \text{ and } \hat{B}_j = y_{.j.} - y_{...}$$

$$\widehat{AB}_{ij} = y_{ij.} - (\hat{\mu} + \hat{A}_i + \hat{B}_j) = y_{ij.} - y_{i..} - y_{.j.} + y_{...}$$

Medication A			
Biofeedback B	no	yes	Total
no	190	186	$y_{.1.} = 188$
yes	188	168	$y_{.2.} = 178$
Total	$y_{1..} = 189$	$y_{2..} = 177$	$y_{...} = 183$

$$\hat{\mu} = 183, \hat{A}_1 = -\hat{A}_2 = 6, \hat{B}_1 = -\hat{B}_2 = 5$$

## *Predicted values of the additive model*

Treatment means:  $(\hat{\mu} + \hat{A}_i + \hat{B}_j)$

	Medication	
	no	yes
Biofeedback		
no	194	182
yes	184	172

$$\widehat{AB}_{11} = \widehat{AB}_{22} = -\widehat{AB}_{12} = -\widehat{AB}_{21} = 4.$$

# Decomposition of Variability

$$SS_{tot} = SS_A + SS_B + SS_{AB} + SS_{res}$$

$$\begin{aligned}SS_{tot} &= \sum \sum \sum (y_{ijk} - y_{...})^2 \\SS_A &= \sum \sum \sum (y_{i..} - y_{...})^2 \\SS_B &= \sum \sum \sum (y_{.j.} - y_{...})^2 \\SS_{AB} &= \sum \sum \sum (y_{ij.} - y_{i..} - y_{.j.} + y_{...})^2 \\SS_{res} &= \text{«Difference»}\end{aligned}$$

degrees of freedom: main effect with  $I$  levels:  $I - 1$  df,  
interaction between two factors with  $I$  and  $J$  levels:  
 $(I - 1)(J - 1)$  df.



## ***Anova table***

Source	SS	df	MS	F	P value
medi	720	1	720	11.52	0.004
bio	500	1	500	8.00	0.012
medi:bio	320	1	320	5.12	0.038
residual	1000	16	62.5		
total	2540	19			

# Treatment effects

	effect size	C.I.
medi without bio:	$y_{21.} - y_{11.} = 186 - 190 = -4$	$(-18, 10)$
medi with bio:	$y_{22.} - y_{12.} = 168 - 188 = -20$	$(-34, -6)$
bio without medi:	$y_{12.} - y_{11.} = 188 - 190 = -2$	$(-16, 12)$
bio with medi:	$y_{22.} - y_{21.} = 168 - 186 = -18$	$(-32, -4)$

(standard error:  $\sqrt{2 \cdot MS_{res} / 5} = 5$ )

# Efficiency of factorials

factorial		
medication		
biofeedback	no	yes
no	6	6
yes	6	6

Two separate studies

medication	
no	yes
6	6

and

biofeedback	
no	yes
6	6

or:

control	medi	bio
8	8	8

## ***More than two factors***

	bio/medi	bio	medi	control
without diet	158	188	186	185
	163	183	191	190
	173	198	196	195
	178	178	181	200
	168	193	176	180
with diet	162	162	164	205
	158	184	190	199
	153	183	169	171
	182	156	165	161
	190	180	177	179

# Treatment means

without diet(C)				with diet (C)			
medication (A)				medication (A)			
bio (B)	no	yes	total	bio (B)	no	yes	total
no	190	186	188	no	183	173	178
yes	188	168	178	yes	173	169	171
total	189	177	183	total	178	171	174.5

# Main effects and interactions

**main effects:** medication (A):  $183.5 - 174 = 9.5$   
biofeedback (B):  $183 - 174.5 = 8.5$   
diet(C):  $183 - 174.5 = 8.5$

## 2-way interactions AB, AC, BC:

	medi (A)		
Bio (B)	no	yes	total
no	186.5	179.5	183
yes	180.5	168.5	174.5
total	183.5	174	178.75

## 3-way interaction ABC

## Model and Anova table

$$Y_{ijkl} = \mu + A_i + B_j + C_k + (AB)_{ij} + (AC)_{ik} + (BC)_{jk} + (ABC)_{ijk} + \epsilon_{ijkl}$$

$$\sum A_i = 0, \dots, \sum_k (ABC)_{ijk} = 0$$

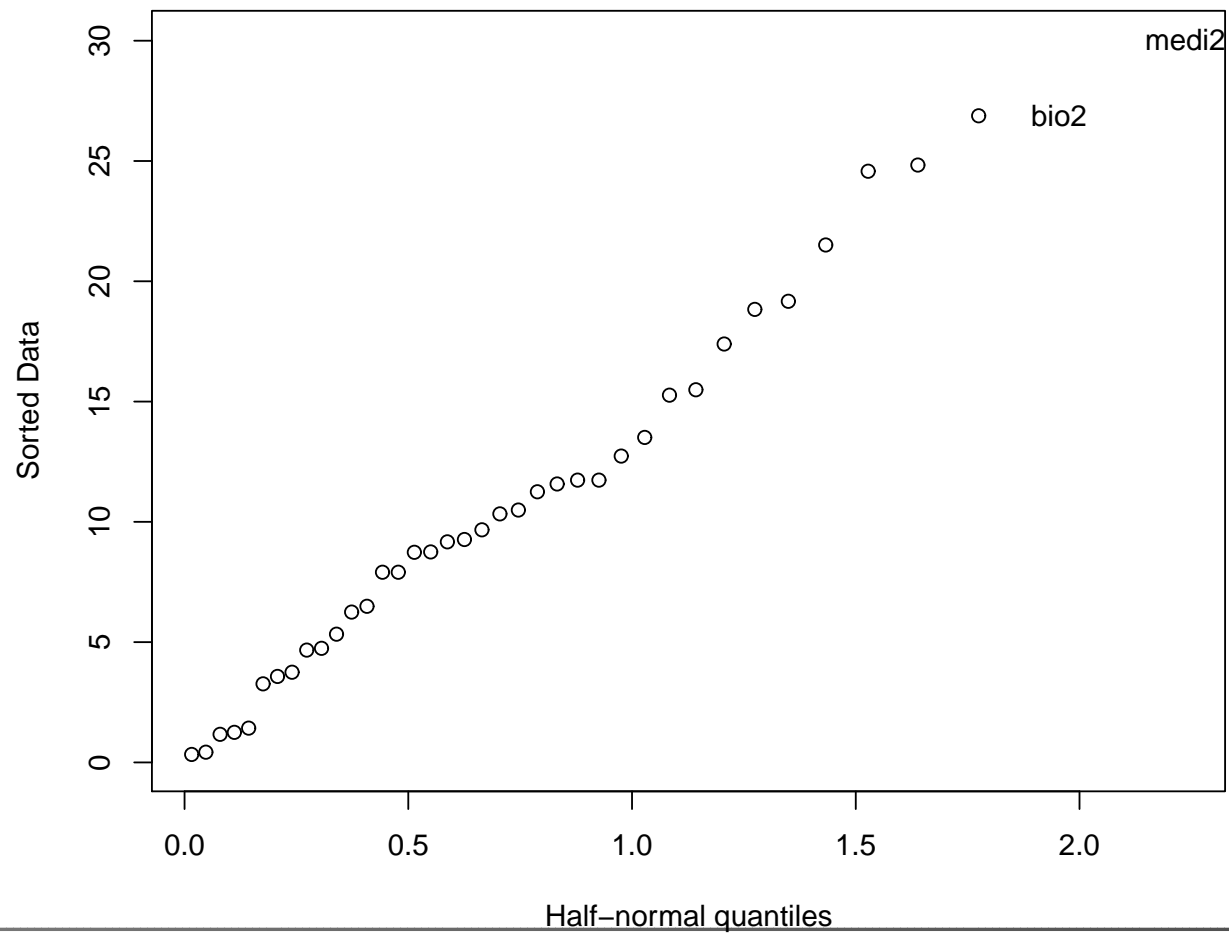
Source	df	MS=SS/df	F
A	$I - 1$		$MS_A / MS_{res}$
B	$J - 1$		$MS_B / MS_{res}$
C	$K - 1$		$MS_C / MS_{res}$
AB	$(I - 1)(J - 1)$		$MS_{AB} / MS_{res}$
AC	$(I - 1)(K - 1)$		$MS_{AC} / MS_{res}$
BC	$(J - 1)(K - 1)$		$MS_{BC} / MS_{res}$
ABC	$(I - 1)(J - 1)(K - 1)$		$MS_{ABC} / MS_{res}$
Residual	«Difference»	$MS_{res} = \hat{\sigma}^2$	
Total	$IJKn - 1$		

## ***Anova table***

Source	SS	df	MS	F	P value
medi	902.5	1	902.5	6.33	0.017
bio	722.5	1	722.5	5.06	0.031
diet	722.5	1	722.5	5.06	0.031
medi:bio	62.5	1	62.5	0.44	0.51
medi:diet	62.5	1	62.5	0.44	0.51
bio:diet	22.5	1	22.5	0.16	0.69
medi:bio:diet	302.5	1	302.5	2.12	0.15
Residual	4566.0	32	142.7		
Total	7363.5	39			



# Half normal plot



# Unbalanced Factorials

uncorrelated estimators:

$$SS_{tot} = SS_A + SS_B + SS_{AB} + \underbrace{SS_{res}}_{SS_C + \dots + SS_{res'}}$$

correlated estimators:

$$SS_{tot} = SS'_A + SS'_B + SS'_{AB} + SS_C + \dots + SS_{res'}$$

**SS Typ I:**  $SS_A$  ignores all other  $SS$

**SS Typ II:**  $SS_A$  takes into account all other main effects, ignores all interactions

**SS Typ III:**  $SS_A$  takes into account all other effects

# Calculation of SS's

by model comparison

For SS Typ I:

model 1: $Y_{ijk} = \mu + \epsilon_{ijk}$	$SS_{e1} = SS_T$
model 2: $Y_{ijk} = \mu + A_i + \epsilon_{ijk}$	$SS_{e2}$
model 3: $Y_{ijk} = \mu + A_i + B_j + \epsilon_{ijk}$	$SS_{e3}$
model 4: $Y_{ijk} = \mu + A_i + B_j + AB_{ij} + \epsilon_{ijk}$	$SS_{e4} = SS_{res}$

# *Rat genotype*

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- Litters of rats are separated from their natural mother and given to another female to raise.
- 2 factors: mother's genotype (A, B, I, J) and litter's genotype (A, B, I, J)
- response: average weight gain of the litter.

# Full model

```
> summary(aov(y~mother*genotype,data=gen))
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)	
mother	3	771.61	257.202	4.7419	0.005869	*
genotype	3	63.63	21.211	0.3911	0.760004	
mother:genotype	9	824.07	91.564	1.6881	0.120053	
Residuals	45	2440.82	54.240			

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```
> summary(aov(y~genotype*mother,data=gen))
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)	
genotype	3	60.16	20.052	0.3697	0.775221	
mother	3	775.08	258.360	4.7632	0.005736	*
genotype:mother	9	824.07	91.564	1.6881	0.120053	
Residuals	45	2440.82	54.240			

# SS Typ II in

```
> drop1(mod1, test="F")
```

```
Model: y ~ mother * genotype
```

	Df	Sum Sq	RSS	AIC	F value	Pr(F)
<none>			2440.8	257.04		
mother:geno	9	824.07	3264.9	256.79	1.6881	0.1201

```
> drop1(mod1, .~., test="F")
```

```
Model: y ~ mother * genotype
```

	Df	Sum Sq	RSS	AIC	F value	Pr(F)
<none>			2440.8	257.04		
mother	3	582.25	3023.1	264.09	3.5782	0.02099 *
geno	3	591.69	3032.5	264.28	3.6362	0.01968 *
mother:geno	9	824.07	3264.9	256.79	1.6881	0.12005

## ***Offer for a 6-year old car***

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- Planned experiment to see whether the offered cash for the same medium-priced car depends on gender or age (young, middle, elderly) of the seller.
- 6 factor combinations with 6 replications each.
- Response variable  $y$  is offer made by a car dealer (in \$ 100)
- Covariable: overall sales volume of the dealer

# *Analysis of Covariance*

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- Covariates can reduce  $MS_{res}$ , thereby increasing power for testing.
- Baseline or pretest values are often used as covariates. A covariate can adjust for differences in characteristics of subjects in the treatment groups.
- It should be related only to the response variable and not to the treatment variables (factors).
- We assume that the covariate will be linearly related to the response and that the relationship will be the same for all levels of the factor (no interaction between covariate and factors).



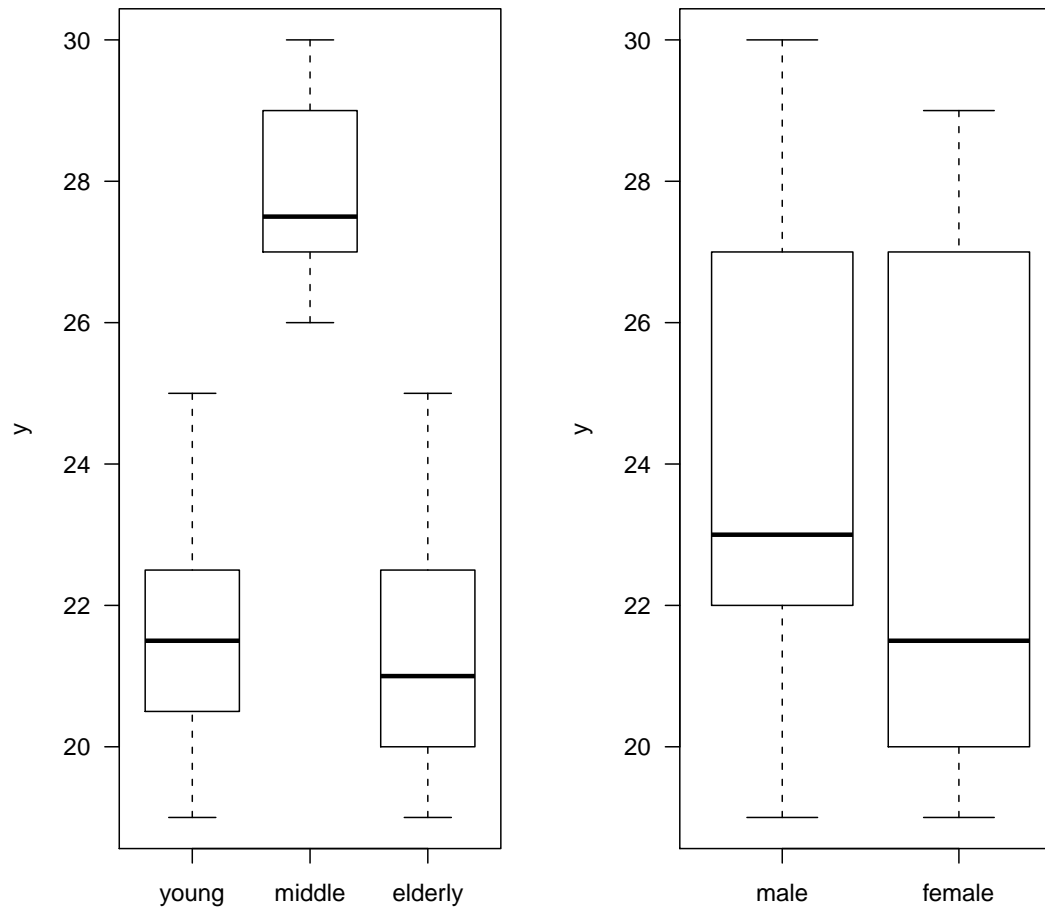
# ***Model for two-way ANCOVA***

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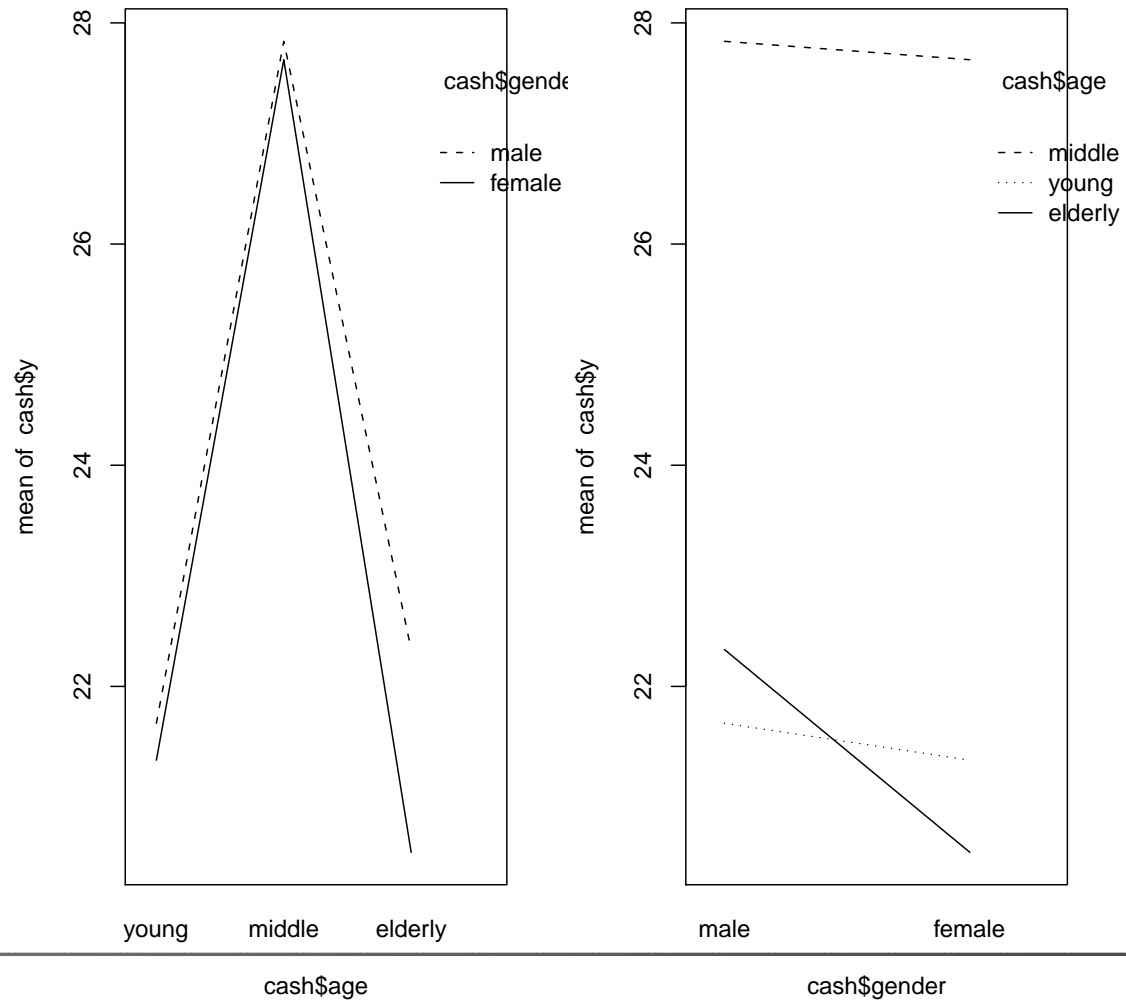
$$Y_{ijk} = \mu + \theta x_{ijk} + A_i + B_j + (AB)_{ij} + \epsilon_{ijk}$$

$$\sum A_i = \sum B_j = \sum (AB)_{ij} = 0, \quad \epsilon_{ijk} \sim \mathcal{N}(0, \sigma^2)$$

# Effect of Age and Gender



# Interaction effect of Age and Gender

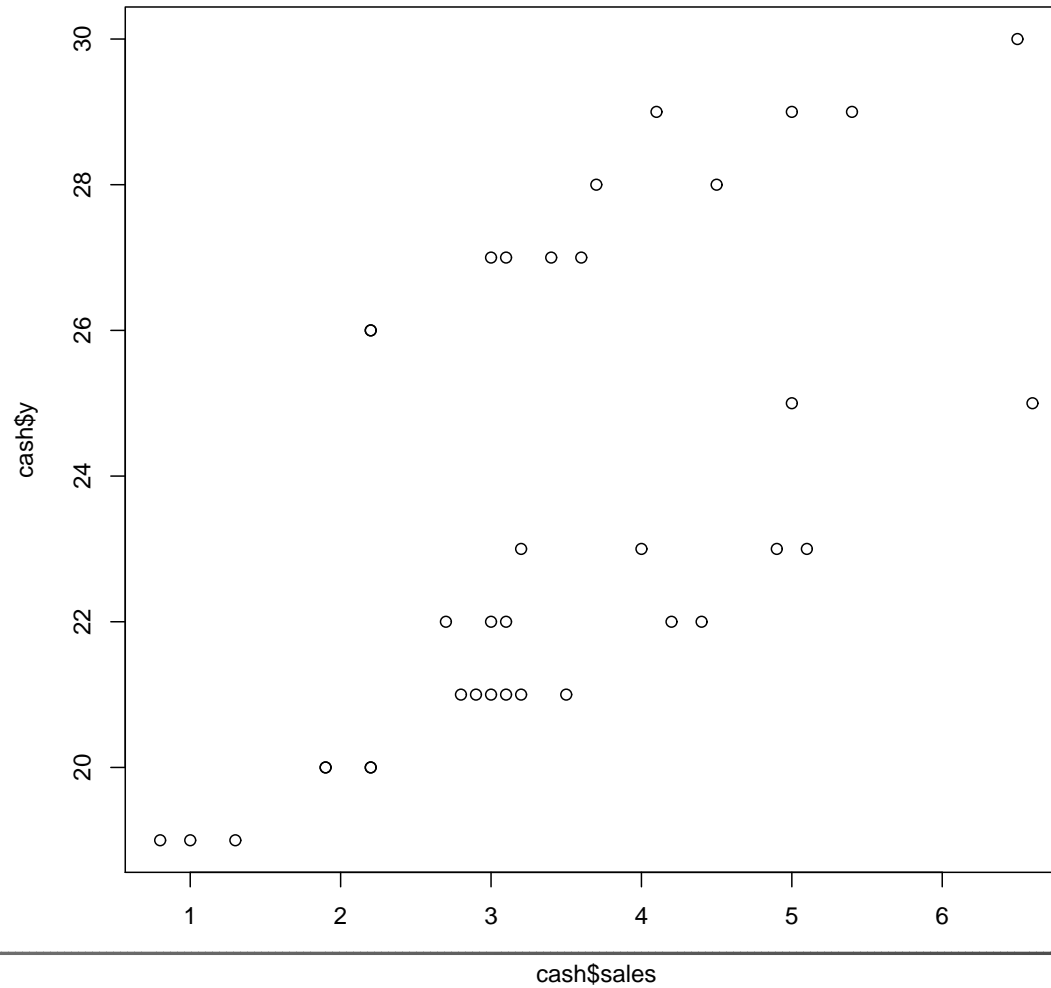


# Two-way Anova

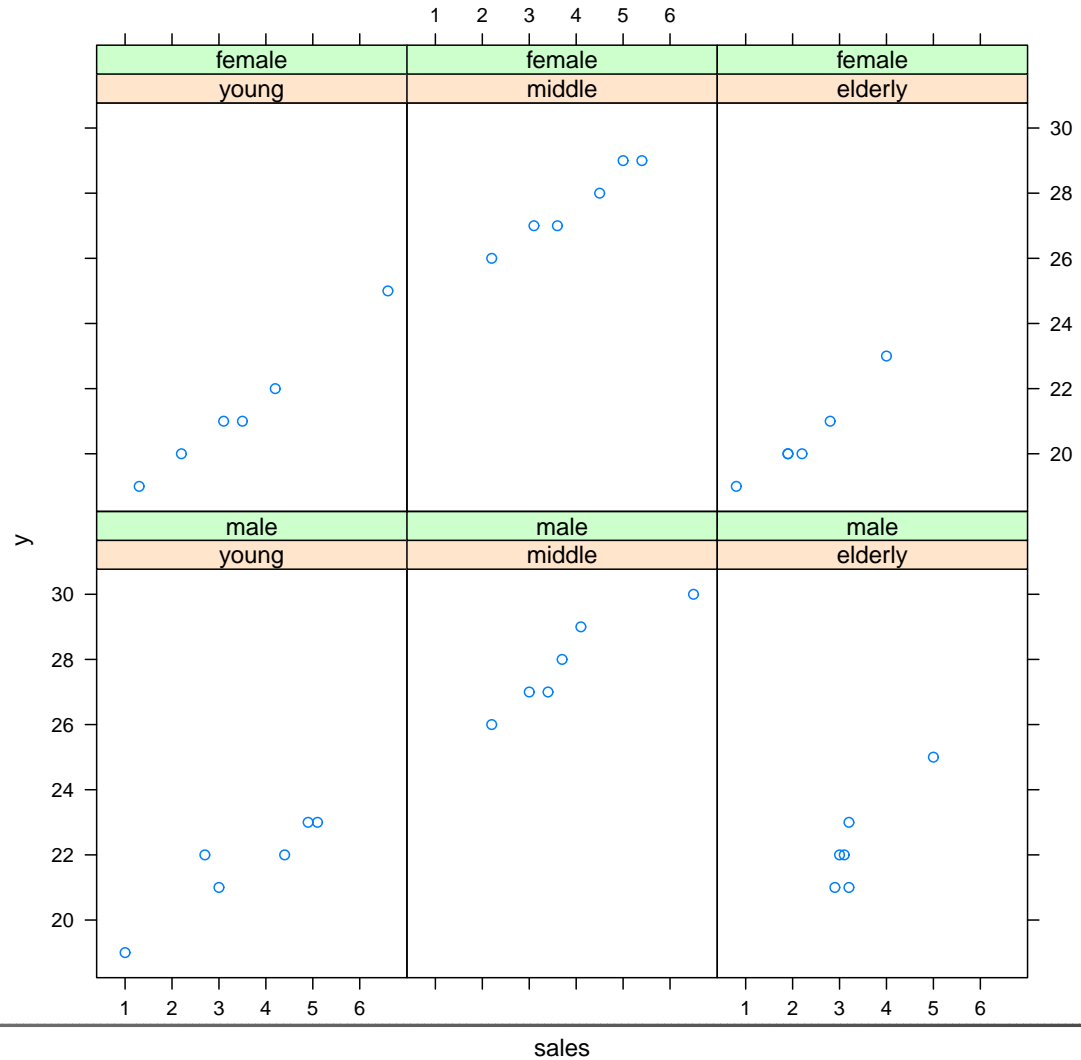
```
> mod1=aov(y~age*gender,data=cash)
> summary(mod1)
```

	Df	Sum Sq	MeanSq	Fvalue	Pr(>F)
age	2	316.72	158.36	66.29	9.79e-12*
gender	1	5.44	5.44	2.28	0.1416
age:gender	2	5.06	2.53	1.06	0.3597
Residuals	30	71.67	2.39		

# Sales and Cash Offer



# Sales and Cash Offer by Group



# Two-way Ancova

```
> mod2=aov(y~sales+age*gender,data=cash)
```

```
> summary(mod2)
```

	Df	SumSq	MeanSq	Fvalue	Pr(>F)
sales	1	157.37	157.37	550.22	< 2e-16**
age	2	231.52	115.76	404.75	< 2e-16**
gender	1	1.51	1.51	5.30	0.02874*
age:gender	2	0.19	0.10	0.34	0.71422
Residuals	29	8.29	0.29		