

Inverse Problems in Statistics

Chapter 4: Some Methods of Estimation

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1 Plug-in Estimators

The plug-in estimator $\tilde{\theta}_n$ of $\theta = Q(F)$ is defined by $Q(F_n)$, where

$$F_n(x) = \frac{1}{n} \sum_{i=1}^n 1_{\{X_i=x\}}$$

is the empirical distribution function, X_i are independent identically F -distributed.

1.1 Wicksell's problem I

We observe a sample from the density g .

The object of interest is the distribution function F , which can be expressed as :

$$F(x) = 1 - \frac{V(x)}{V(0)}$$

where

$$V(x) = \int_x^\infty (z-x)^{-\frac{1}{2}} g(z) dz$$

Plug-in estimator for V :

$$\tilde{V}_n(x) = \frac{1}{n} \sum_{i=1}^n (Z_i - x)^{-\frac{1}{2}} 1_{(x,\infty)}(Z_i)$$

2 Isotonic inverse estimators

With isotonic inverse estimators we can correct the nonmonotonicity.

Method 1: first the plug-in estimator is computed and afterwards the estimator is projected on the set of monotone functions.

Method 2: project the nonparametric estimator of the sampling distribution on the set of possible sampling distributions, then the inversion of this projected estimator yields a distribution function as estimator.

Lemma 2.1 *Let $0 < M < \infty$ and ϕ some a.e. continuous nonnegative integrable function on $[0, M]$. Denote by Φ the primitive*

$$\Phi(x) = \int_0^x \phi(y)dy$$

of ϕ on $[0, M]$. Let Φ^ be the least concave majorant of Φ on $[0, M]$, and ϕ^* its right derivative. If ϕ^* is bounded, then, for each bounded decreasing right continuous function σ on $[0, M]$,*

$$\int_0^M (\sigma^2(x) - 2\sigma(x)\phi(x))dx \geq \int_0^M (\phi^*(x)^2 - 2\phi^*(x)\phi(x))dx$$

Consequently, if ϕ is square integrable,

$$\int_0^M (\sigma(x) - \phi(x))^2 dx \geq \int_0^M (\phi^*(x) - \phi(x))^2 dx$$

for all such σ .

2.1 Wicksell's problem II

$\widetilde{U}_n(x)$ is the primitive of \widetilde{V}_n ,

$$\widetilde{U}_n(x) = 2 \int_{z=0}^{\infty} \sqrt{z} dG_n(z) - 2 \int_x^{\infty} \sqrt{z-x} dG_n(z)$$

\widetilde{U}_n^* is the concave majorant of \widetilde{U}_n , and the right continuous evaluation of the derivative of \widetilde{U}_n^* is the projected inverse estimator for V (the isotonic inverse estimator for V).

3 Maximum Likelihood estimators

The likelihood function is defined by

$$l(g) = \sum_{i=1}^n \log g(z_i)$$

where z_1, \dots, z_n is a given realized sample from density g .

The MLE is defined as the maximizer of $l(g)$ over some class of densities.

The MLE is also an isotonic inverse estimator. We can have two kinds of complications with the MLE: the next two subsections show cases where MLE isn't well defined and the last subsection shows the case of a not uniquely defined MLE.

3.1 Wicksell's problem III

Relation between the density g of the observable data Z and the distribution function V :

$$g(z) = -\frac{1}{\pi} \int_z^\infty \frac{dV(x)}{\sqrt{x-z}}$$

Restriction for V :

$$\int_0^\infty \sqrt{x} dV(x) = -\frac{\pi}{2}$$

The likelihood function can be made infinity, so we change it as following to solve this problem and be able to maximize it:

$$l(V) = \frac{1}{n} \sum_{i=0}^{n-1} \log \left(\sum_{j=i+1}^n \frac{V(z_i) - V(z_j)}{\sqrt{z_j - z_i}} \right)$$

New restriction for V :

$$\sum_{i=1}^n \sqrt{z_i} (V(z_{i-1}) - V(z_i)) = \frac{\pi}{2}$$

3.2 Double censoring

Density of the observable data Z :

$$g(y, \delta, \gamma) = (F(y)h_L(y))^\delta ((H_L(y) - H_R(y))f(y))^\gamma ((1 - F(y))h_R(y))^{1-\gamma-\delta}$$

In this case the likelihood function isn't well defined, a way out of the problem is to use the empirical likelihood \tilde{l} :

$$\tilde{l}(F) = \sum_{i=1}^n \delta_i \log F(y_i) + \gamma_i \log F(\{y_i\}) + (1 - \delta_i - \gamma_i) \log(1 - F(y_i))$$

3.3 Interval censoring case 1 problem

Density g of the observable data Z :

$$g(t, \delta) = F(t)^\delta (1 - F(t))^{1-\delta} h(t)$$

The likelihood function

$$l(F) = \sum_{i=1}^n \delta_i \log(\hat{F}_n(t_i)) + (1 - \delta_i) \log(1 - \hat{F}_n(t_i))$$

has multiple maximizers, so we restrict the class of distribution functions to those that are constant between successive time points to have a unique MLE.