

Applied Time Series Analysis

SS 2013 – Week 02

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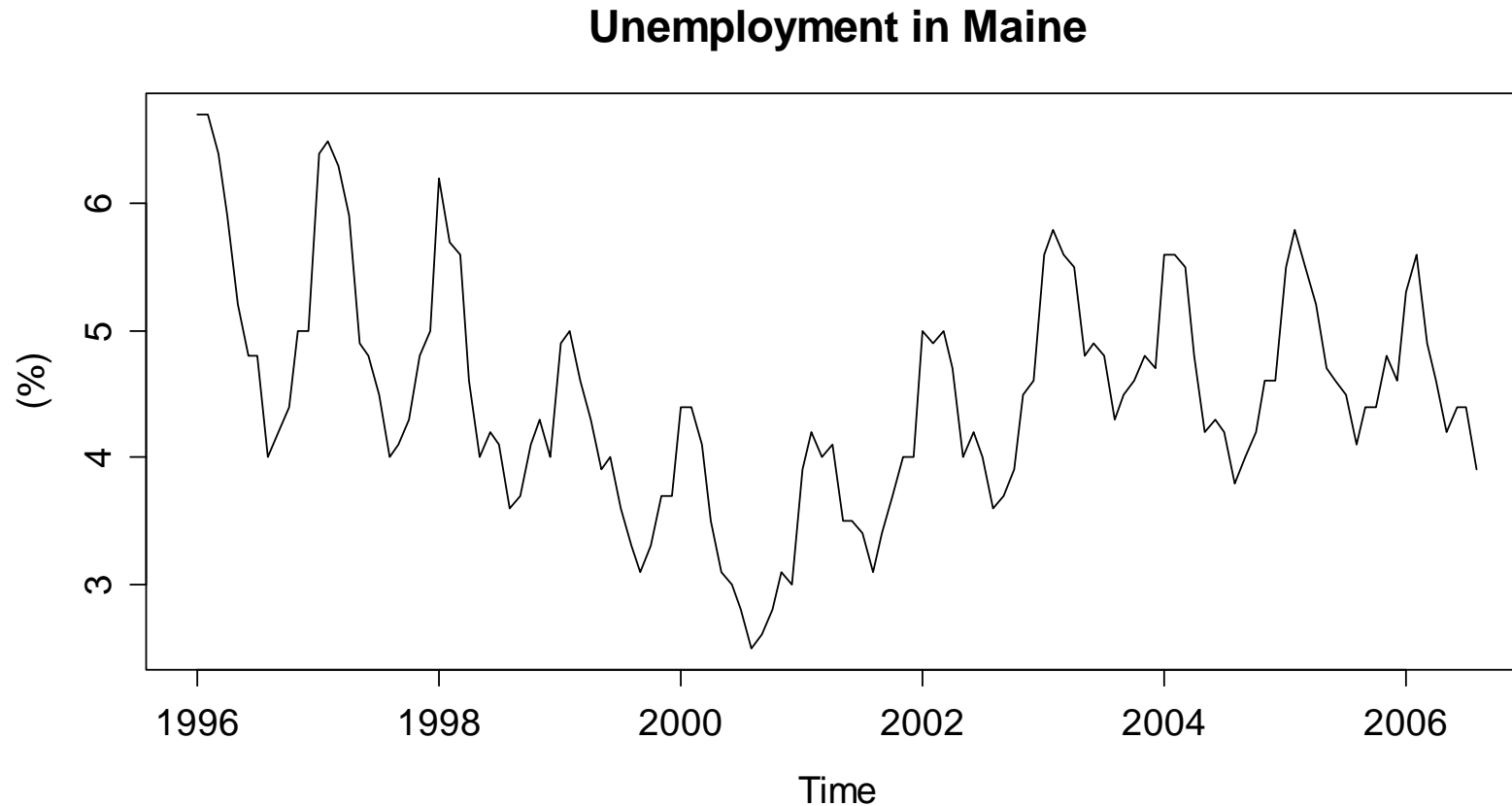
ETH Zürich, February 25, 2013

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Visualization: Time Series Plot

```
> plot(tsd, ylab="(%)", main="Unemployment in Maine")
```



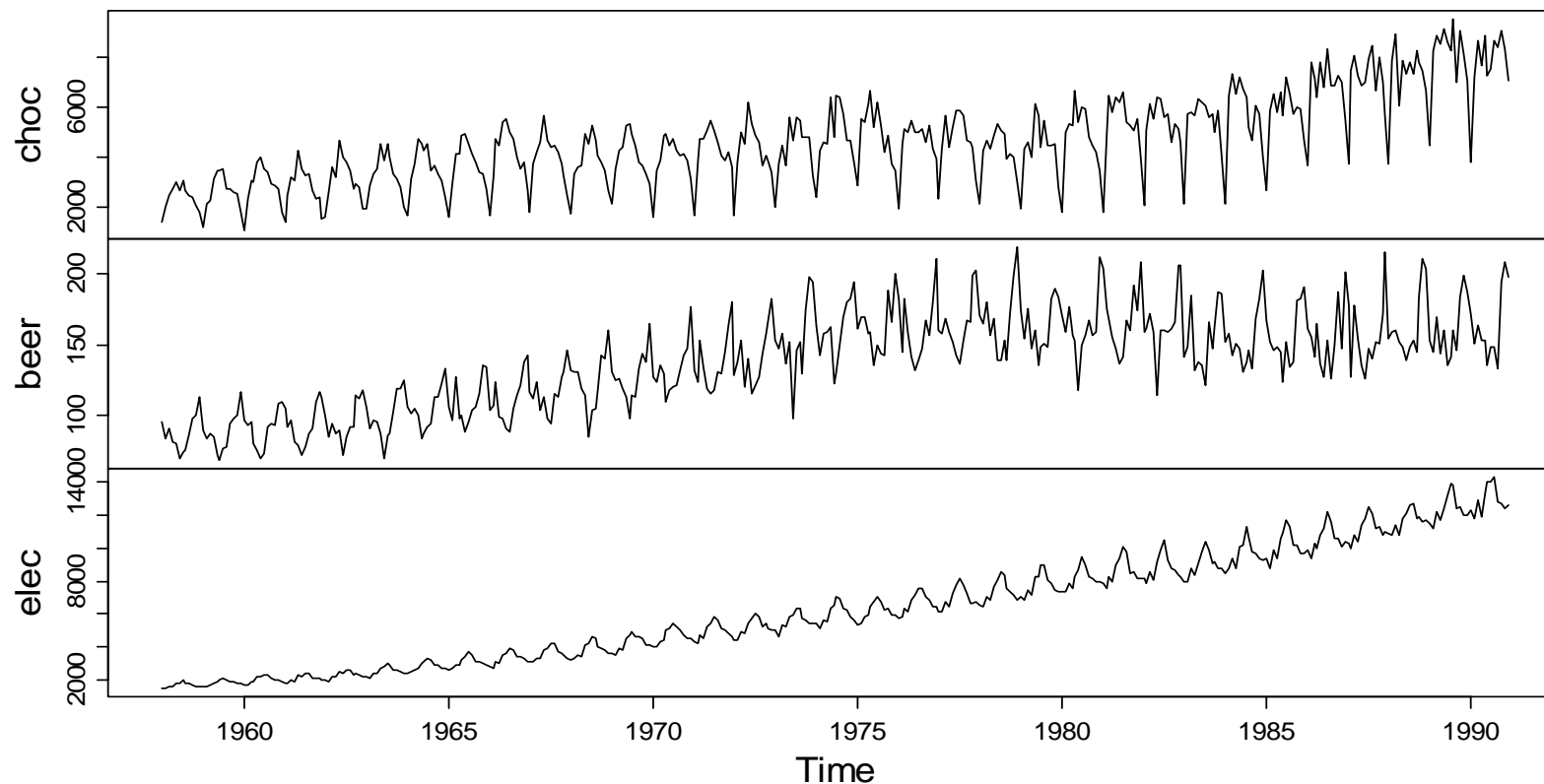
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Multiple Time Series Plots

```
> plot(tsd, main="Chocolate, Beer & Electricity")
```

Chocolate, Beer & Electricity



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Only One or Multiple Frames?

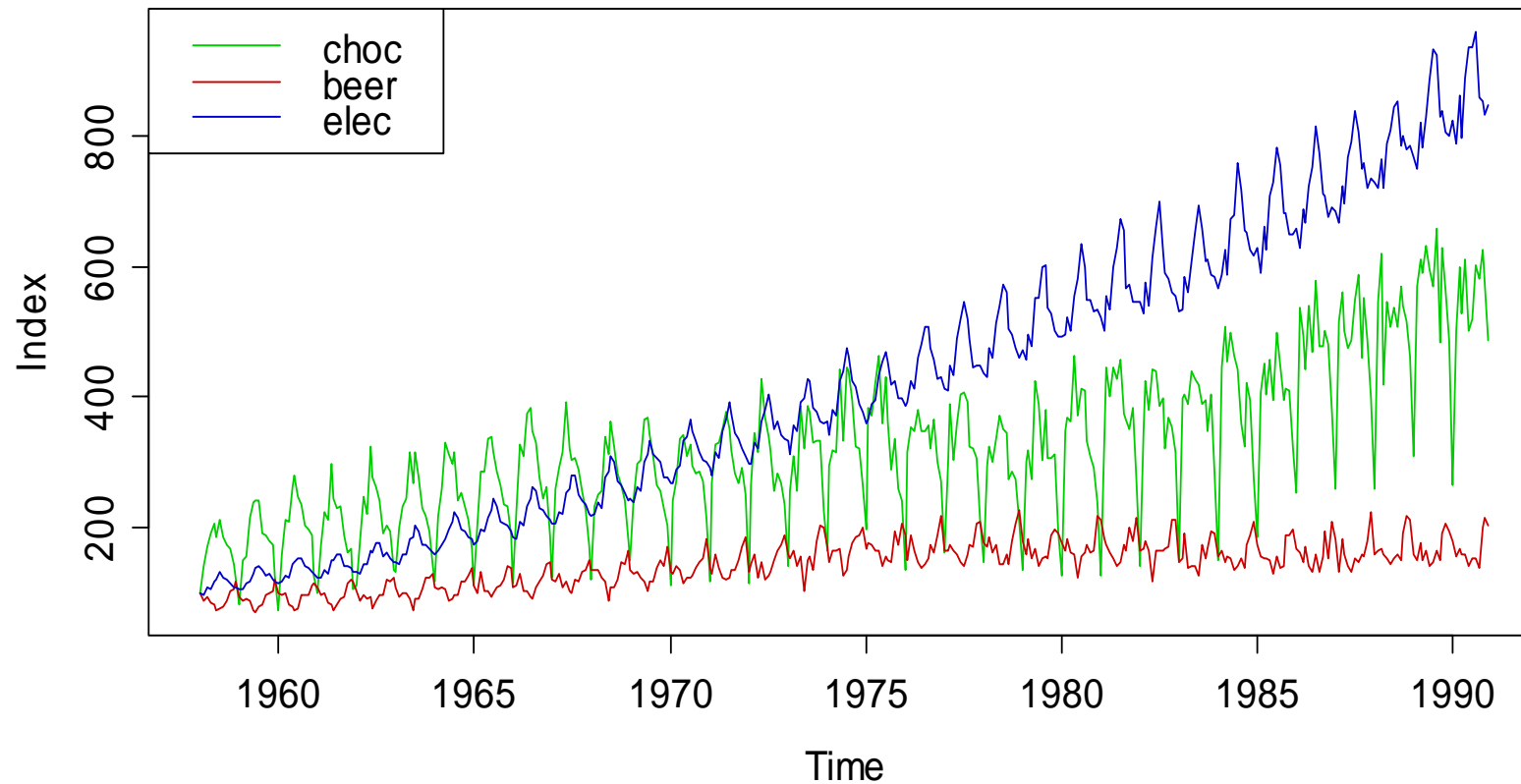
- Due to different scale/units it is often impossible to directly plot multiple time series in one single frame. Also, multiple frames are convenient for visualizing the series.
- If the relative development of multiple series is of interest, then we can (manually) index the series and (manually) plot them into one single frame.
- This clearly shows the magnitudes for trend and seasonality. However, the original units are lost.
- For details on how indexing is done, see the scriptum.

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Multiple Time Series Plots

Indexed Chocolate, Beer & Electricity



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Transformations

For strictly stationary time series, we have: $X_t \sim F$

We did not specify the distribution F and there is no restriction to it. However, many popular time series models are based on:

- 1) *Gaussian distribution*
- 2) *linear relations between the variables*

If the data show different behaviour, we can often improve the situation by transforming x_1, \dots, x_n to $g(x_1), \dots, g(x_n)$. The most popular and practically relevant transformation is:

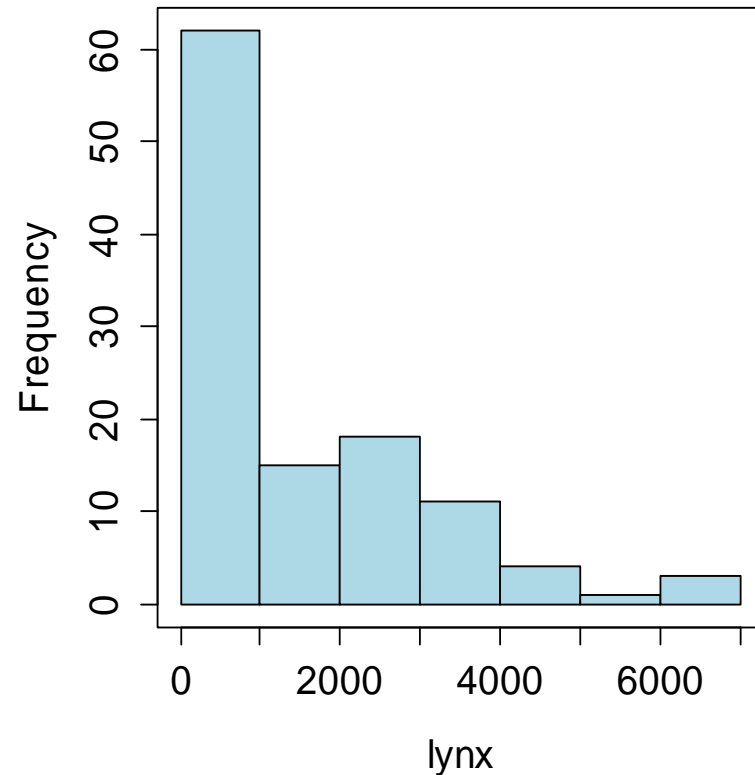
$$g(\cdot) = \log(\cdot)$$

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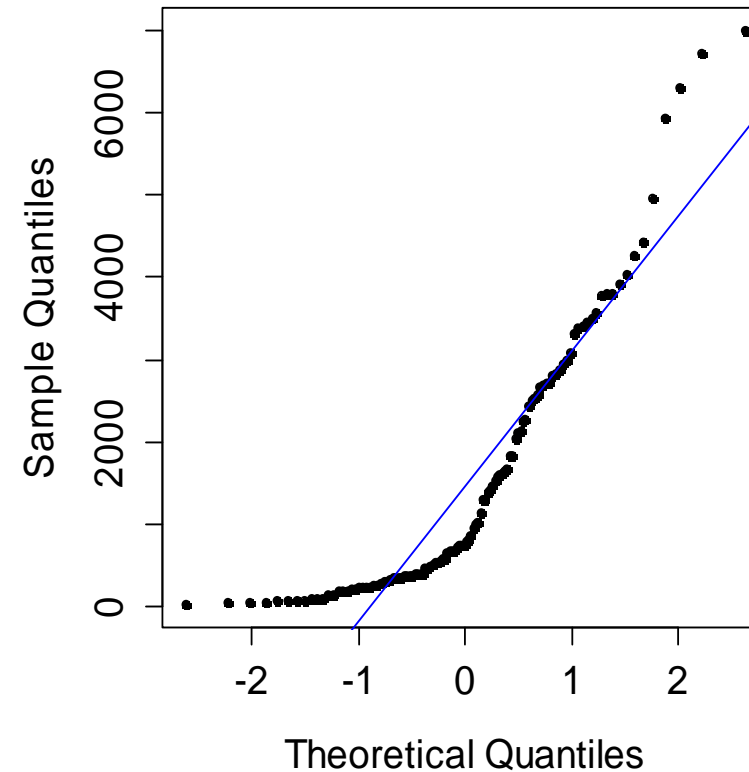
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Transformations: Lynx Data

Histogram of lynx



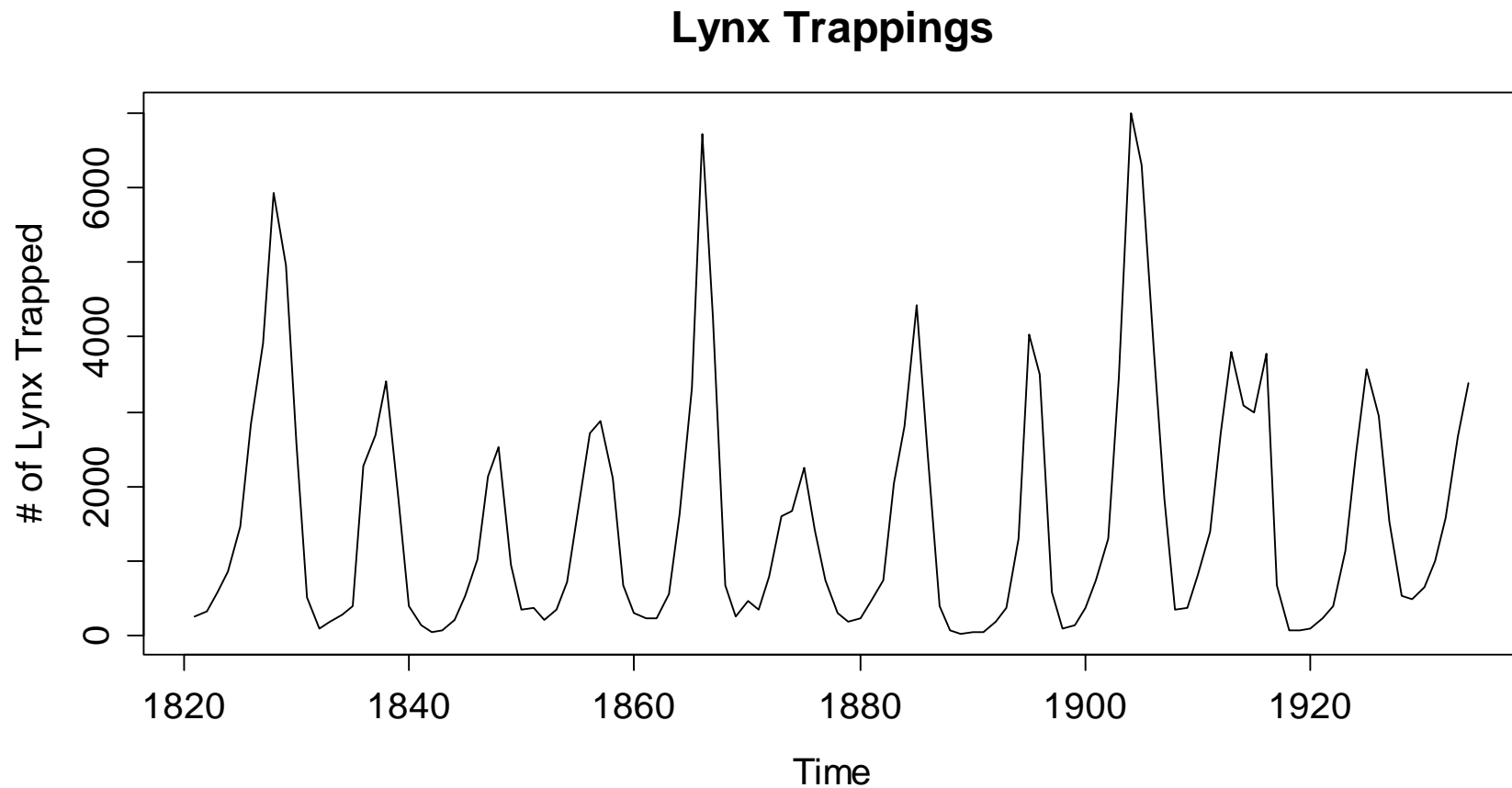
Normal Q-Q Plot



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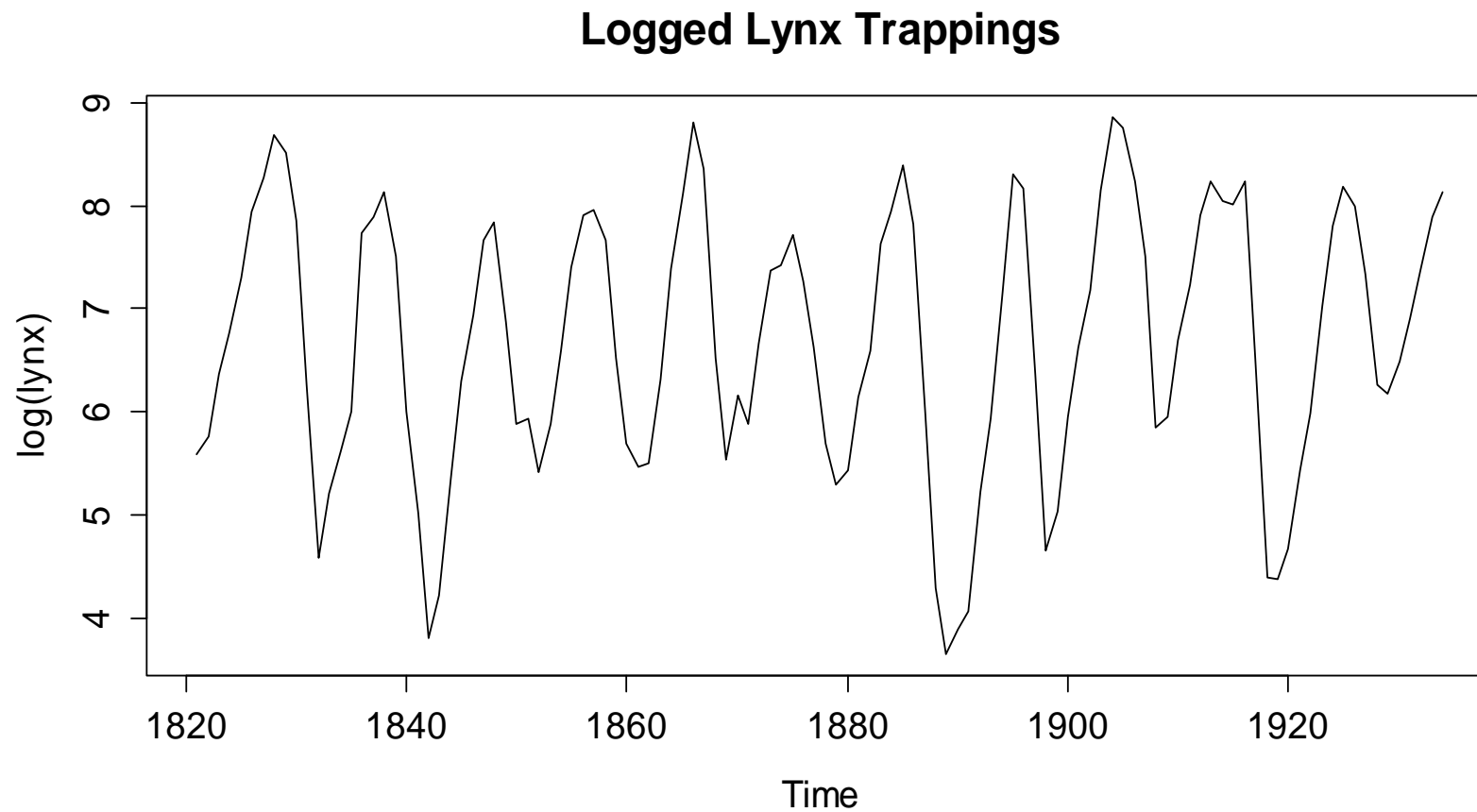
Transformations: Lynx Data



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Transformations: Lynx Data



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Decomposition

Stationarity is key for statistical learning, but real data often have trend/seasonality, and are non-stationary. We can (often) deal with that using the simple additive decomposition model:

$$X_t = m_t + s_t + R_t$$

= trend + seasonal effect + stationary remainder

The goal is to find a remainder term R_t , as a sequence of correlated random variables with mean zero, i.e. a stationary ts.

We can employ:

- 1) *taking differences (=differencing)*
- 2) *smoothing approaches (= filtering)*
- 3) *parametric models (= curve fitting)*

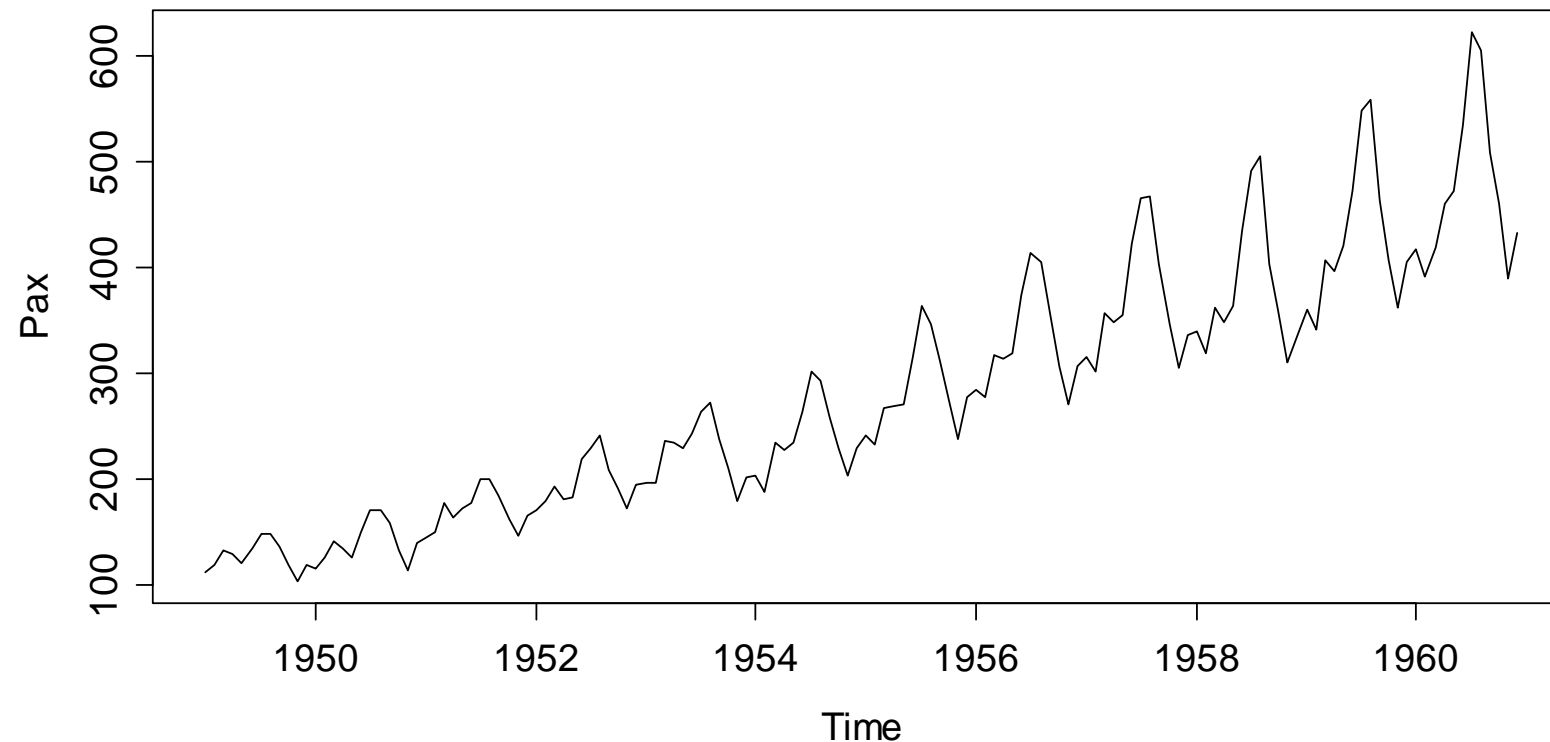
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Multiplicative Decomposition

$X_t = m_t + s_t + R_t$ is not always a good model:

Passenger Bookings

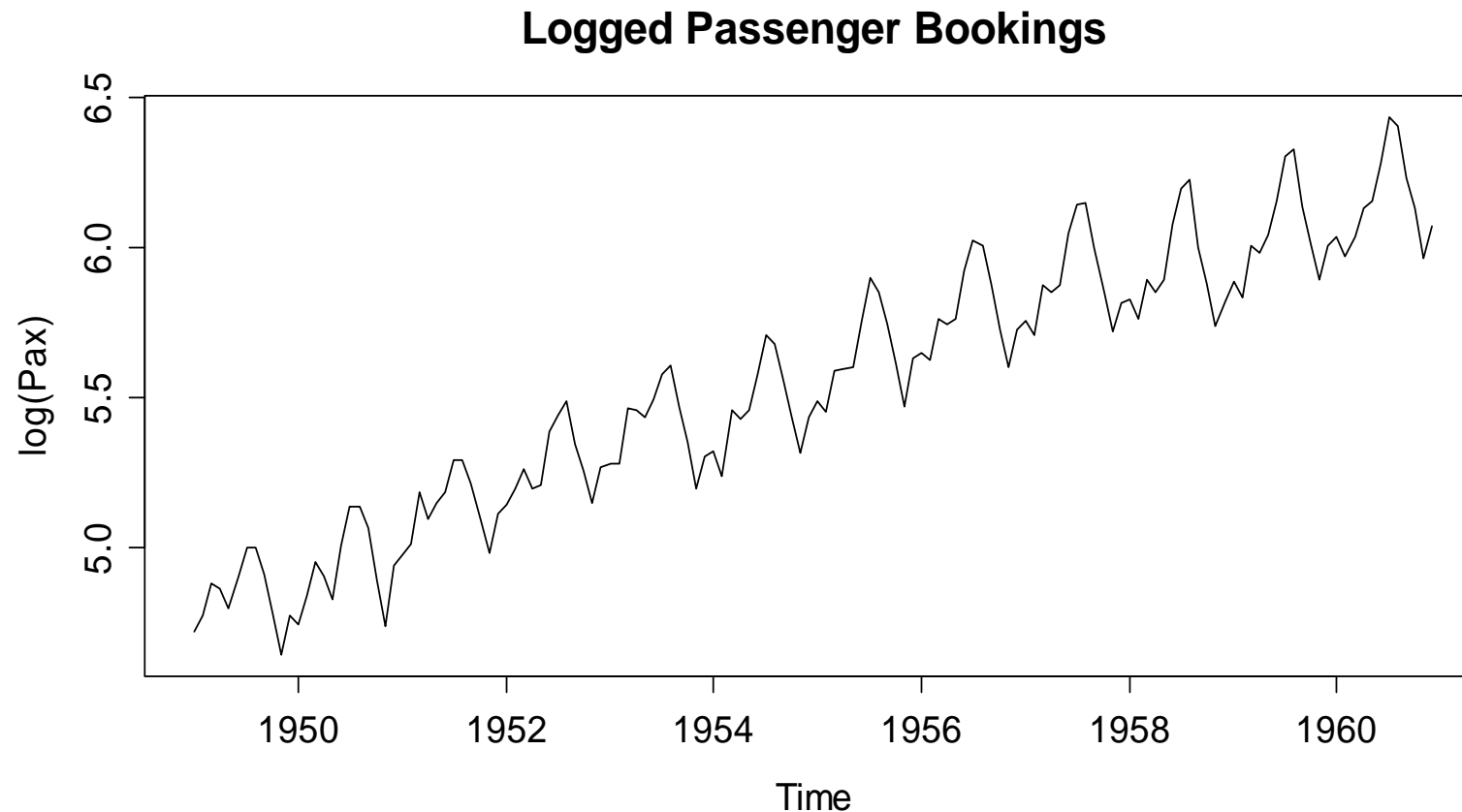


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Multiplicative Decomposition

Better: $X_t = m_t \cdot s_t \cdot R_t$, respectively $\log(X_t) = m'_t + s'_t + R'_t$



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Differencing: Theory

→ [see blackboard...](#)

Summary:

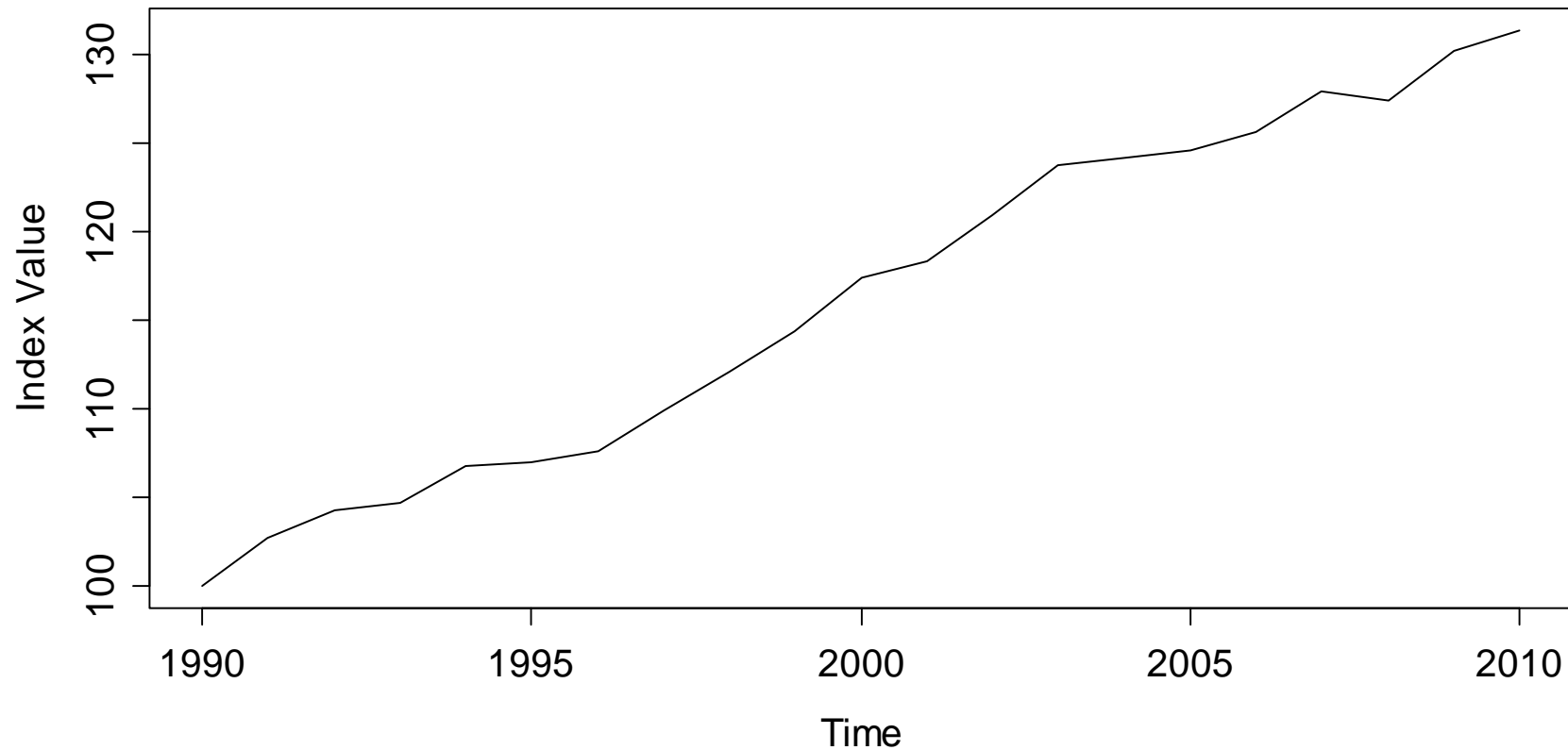
- Differencing means analyzing the observation-to-observation changes in the series, but no longer the original.
- This may (or may not) remove trend/seasonality, but does not yield estimates for m_t and s_t , and not even for R_t .
- Differencing changes the dependency in the series, i.e it artificially creates new correlations.

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Differencing: Example

Swiss Traffic Index

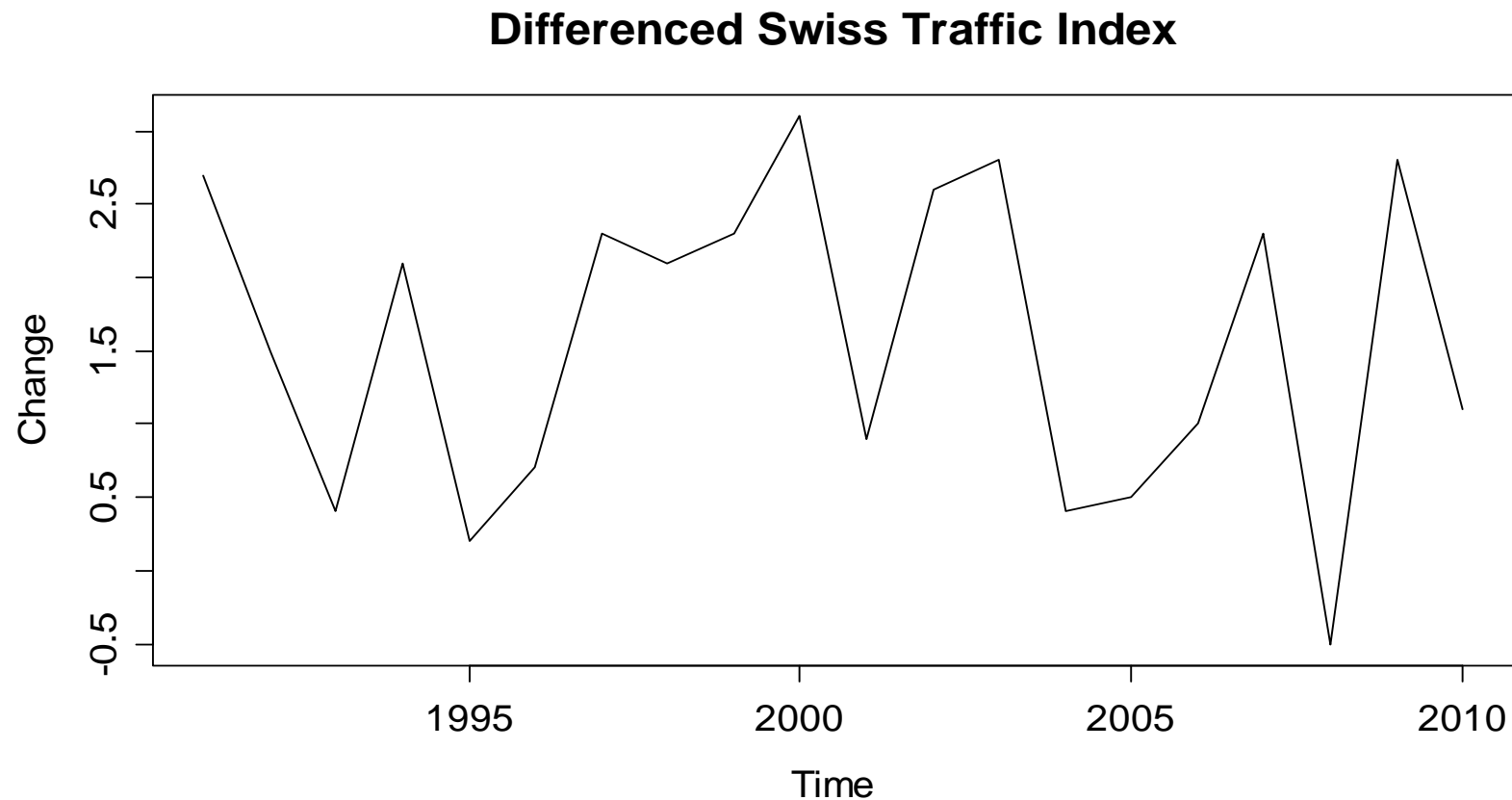


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Differencing: Example

```
> plot(diff(SwissTraffic), main=...)
```



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Differencing: Further Remarks

- If log-transformed series are differenced (i.e. the SMI series), we are considering (an approximation to) the relative changes.

$$Y_t = \log(X_t) - \log(X_{t-1}) = \log\left(\frac{X_t}{X_{t-1}}\right) = \log\left(\frac{X_t - X_{t-1}}{X_{t-1}} + 1\right) \approx \frac{X_t - X_{t-1}}{X_{t-1}}$$

- The backshift operator “go back one step” allows for convenient notation for all differencing operations.

Backshift operator: $B(X_t) = X_{t-1}$

Differencing: $Y_t = (1 - B)X_t = X_t - X_{t-1}$

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Higher-Order Differencing

The “normal” differencing from above managed to remove any linear trend from the data. In case of polynomial trend, that is no longer true. But we can take higher-order differences:

$$\begin{aligned}X_t &= \alpha + \beta_1 t + \beta_2 t^2 + R_t, \quad R_t \text{ stationary} \\Y_t &= (1-B)^2 X_t \\&= (X_t - X_{t-1}) - (X_{t-1} - X_{t-2}) \\&= R_t - 2R_{t-1} + R_{t-2} + 2\beta_2\end{aligned}$$

A quadratic trend can be removed by taking second-order differences. However, what we obtain is not an estimate of the remainder term R_t , but something that is much more complicated.

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Removing Seasonal Effects

Time series with seasonal effects can be made stationary through differencing by comparing to the previous periods' value.

$$Y_t = (1 - B^p)X_t = X_t - X_{t-p}$$

- Here, p is the frequency of the series.
- A potential trend which is exactly linear will be removed by the above form of seasonal differencing.
- In practice, trends are rarely linear but slowly varying: $m_t \approx m_{t-1}$
However, here we compare m_t with m_{t-p} , which means that seasonal differencing often fails to remove trends completely.

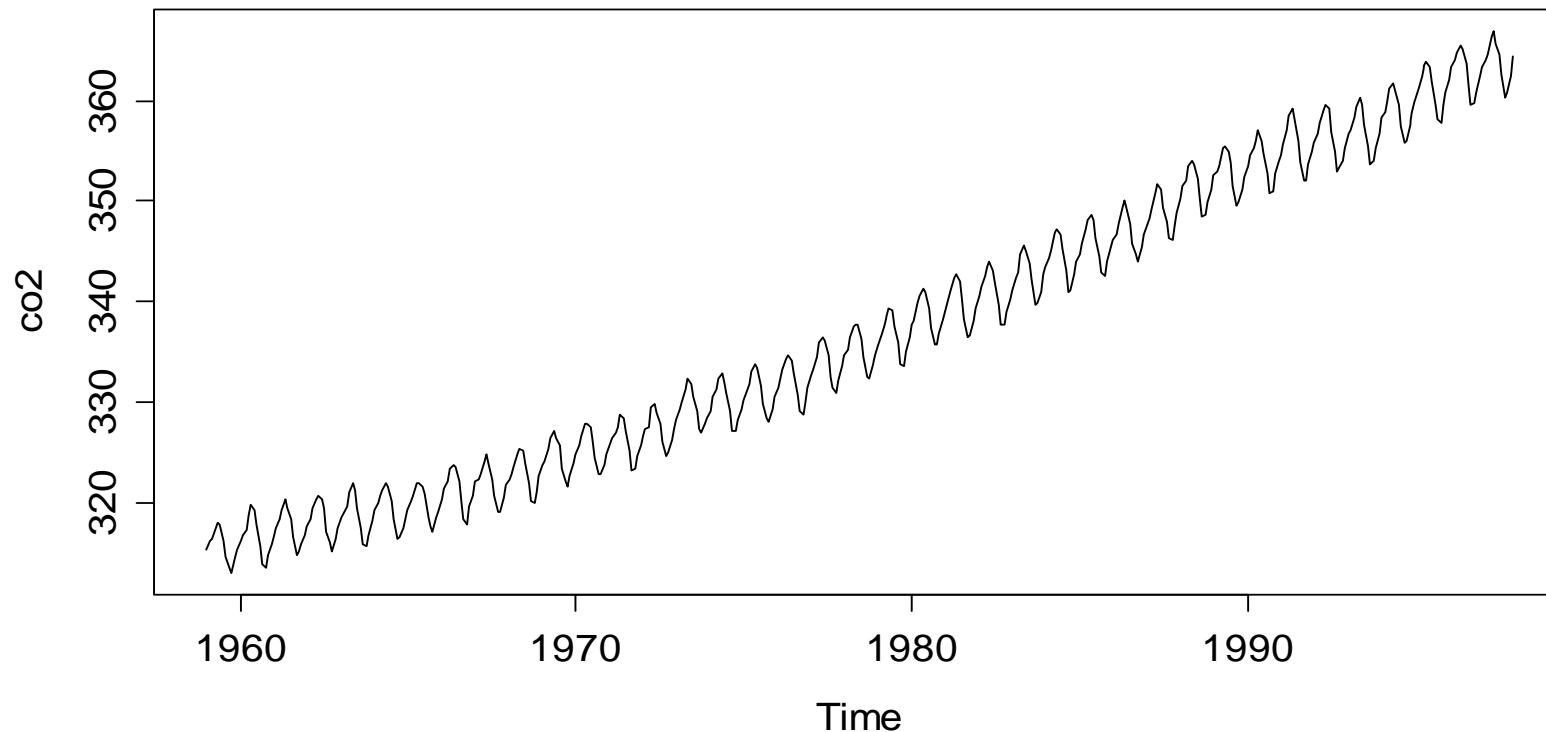
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Seasonal Differencing: Example

```
> data(co2); plot(co2, main=...)
```

Mauna Loa CO2 Concentrations



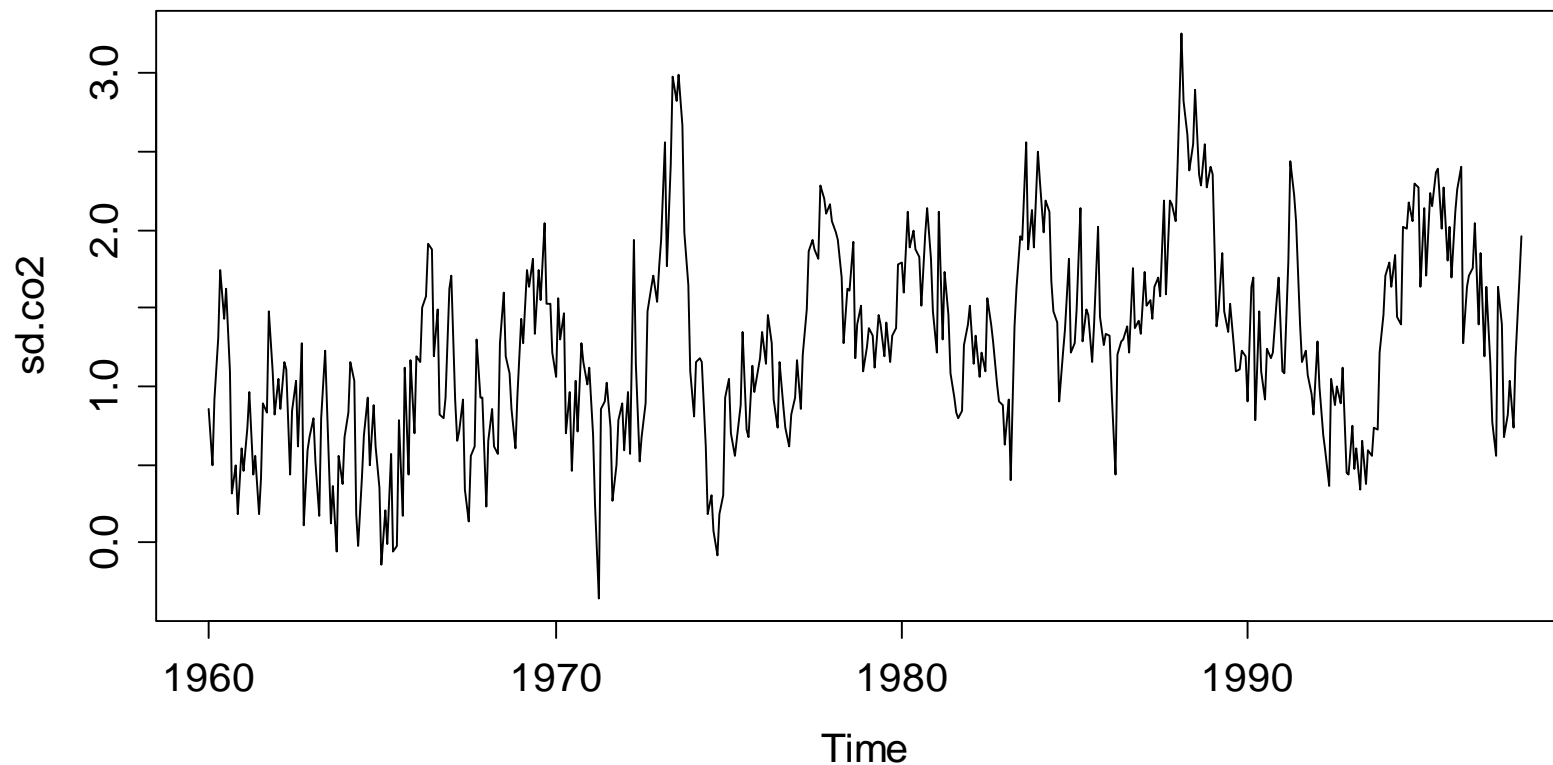
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Seasonal Differencing: Example

```
> sd.co2 <- diff(co2, lag=12)
```

Differenced Mauna Loa Data (p=12)



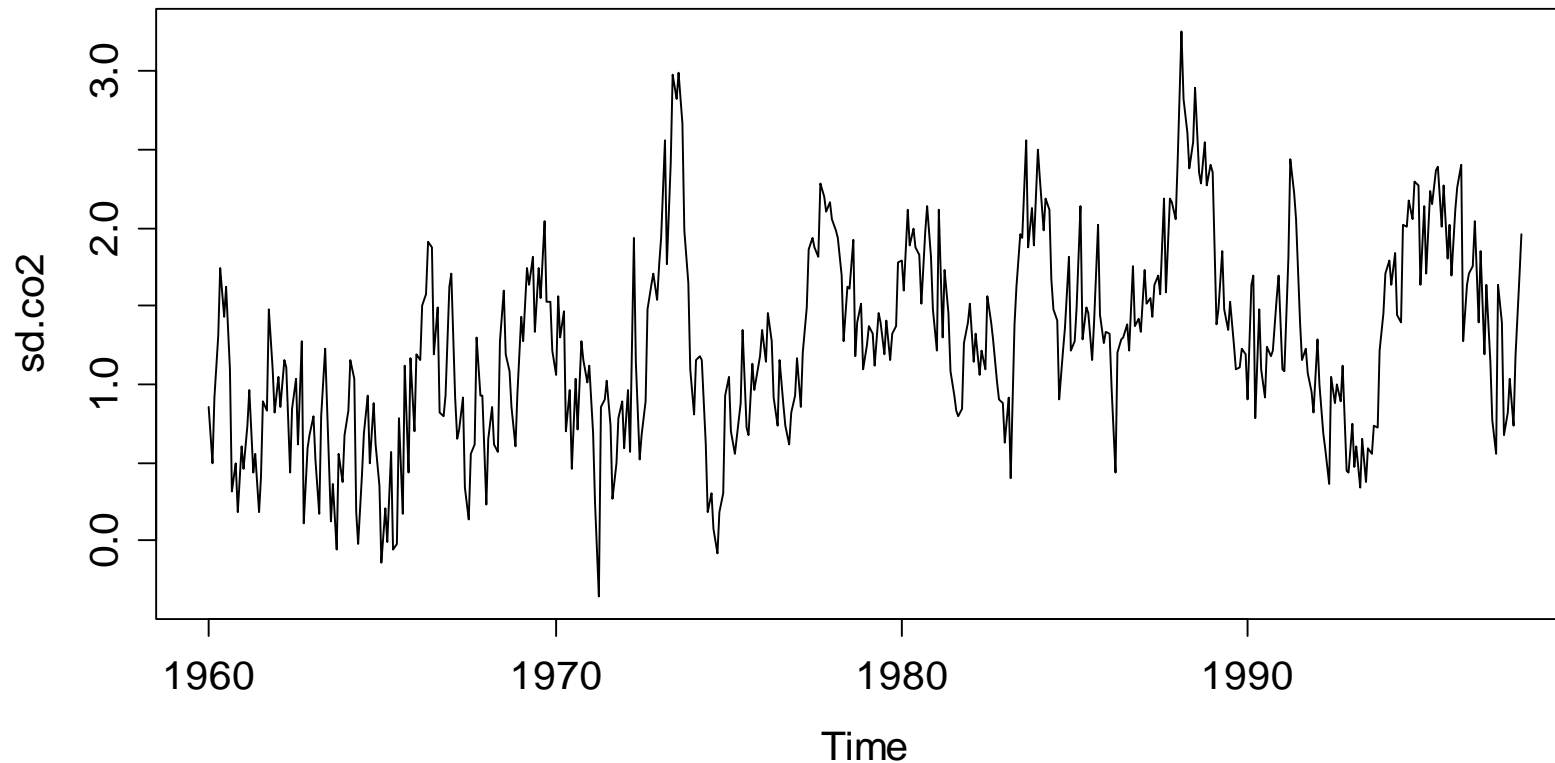
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Seasonal Differencing: Example

```
> sd.co2 <- diff(co2, lag=12)
```

Differenced Mauna Loa Data (p=12)



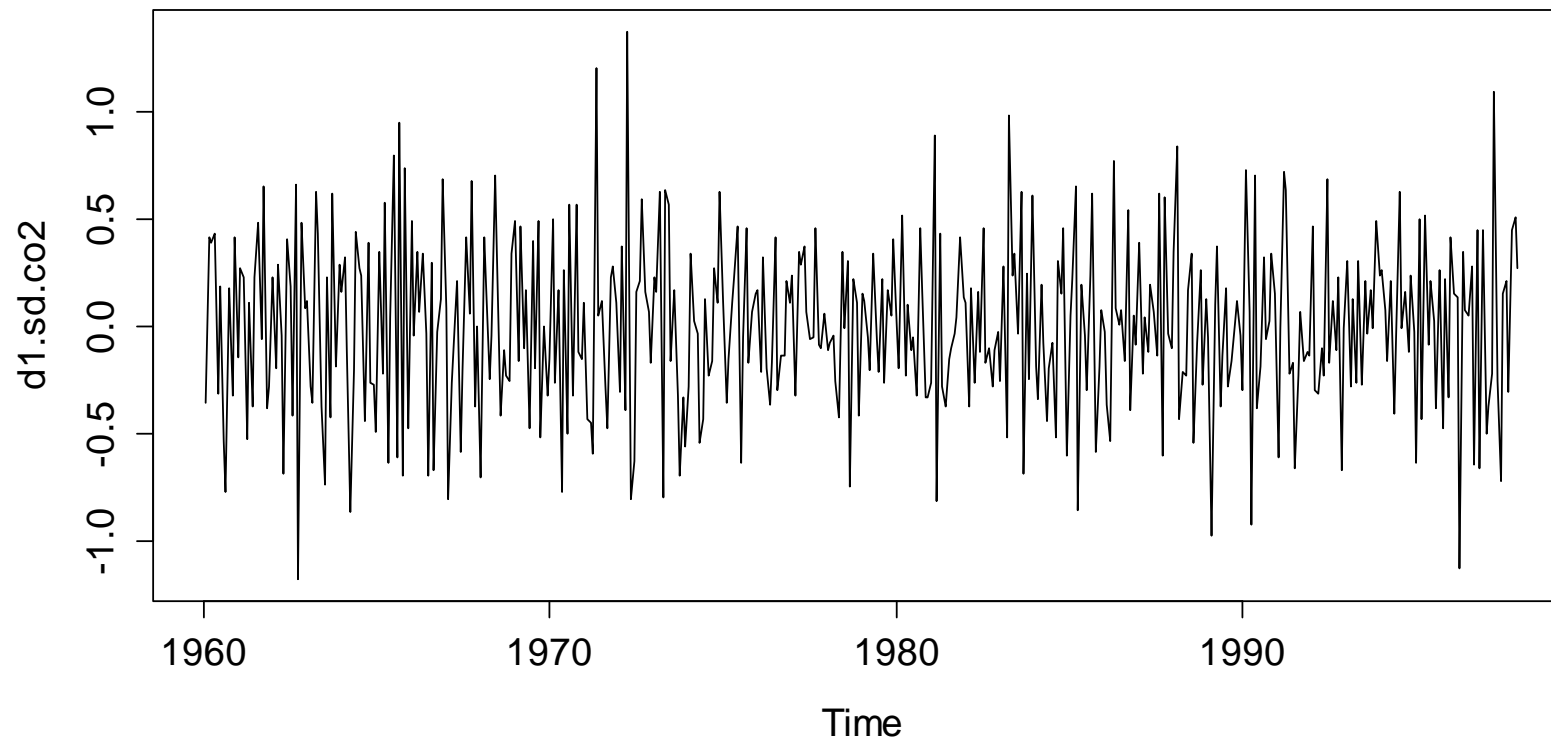
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Seasonal Differencing: Example

This is: $Z_t = (1 - B)Y_t = (1 - B)(1 - B^{12})X_t$

Twice Differenced Mauna Loa Data (p=12, p=1)



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Differencing: Remarks

Some advantages and disadvantages:

- + trend and seasonal effect can be removed
- + procedure is very quick and very simple to implement

- \hat{m}_t , \hat{s}_t and \hat{R}_t are not known, and cannot be visualised
- resulting time series will be shorter than the original
- differencing leads to strong artificial dependencies
- extrapolation of \hat{m}_t , \hat{s}_t is not possible

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Smoothing, Filtering: Part 1

In the **absence of a seasonal effect**, the trend of a non-stationary time series can be determined by applying any **additive, linear filter**. We obtain a new time series \hat{m}_t , representing the trend:

$$\hat{m}_t = \sum_{i=-p}^q a_i X_{t+i}$$

- the window, defined by p and q , can or can't be symmetric
- the weights, given by a_i , can or can't be uniformly distributed
- other smoothing procedures can be applied, too.

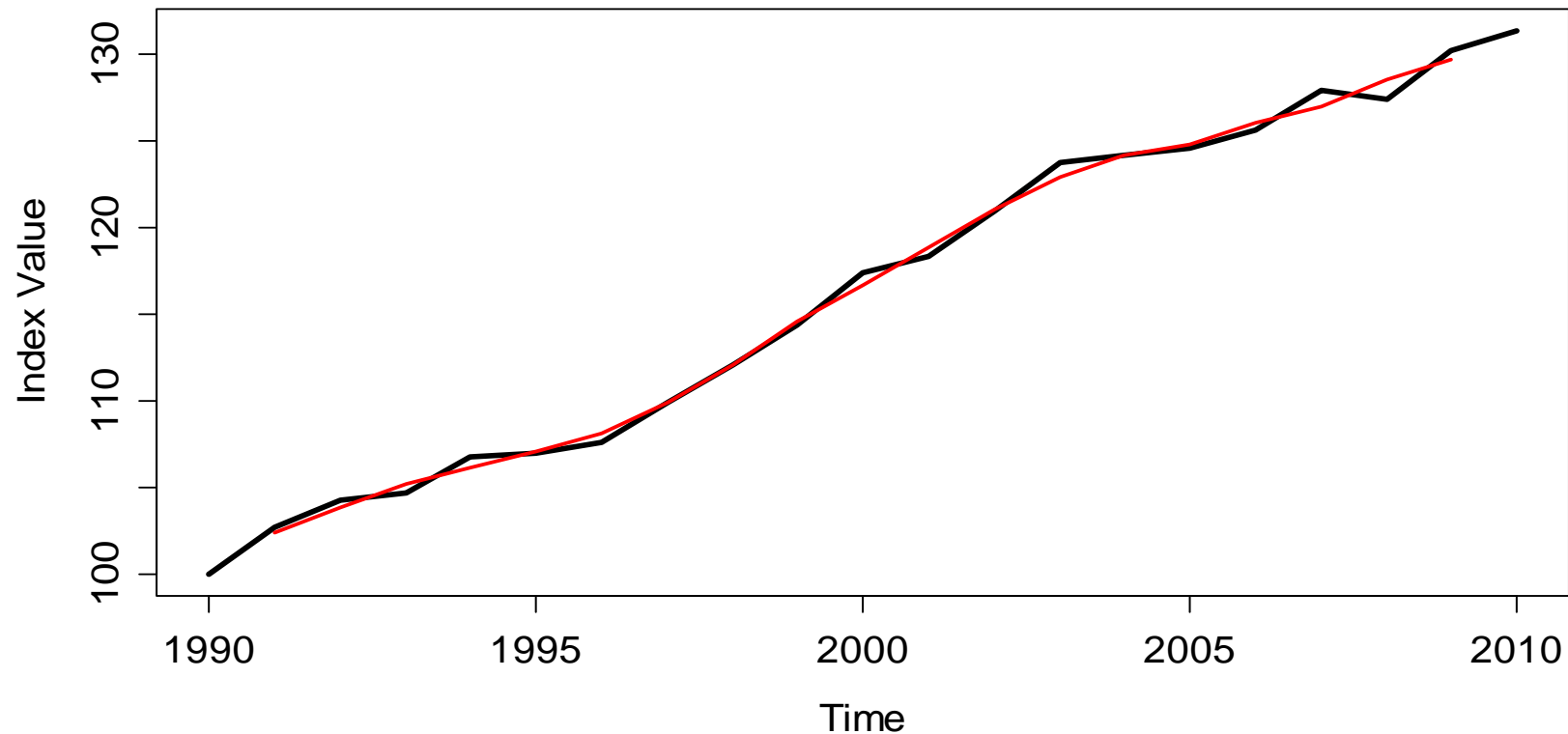
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Trend Estimation with the Running Mean

```
> trd <- filter(SwissTraffic, filter=c(1,1,1)/3)
```

Swiss Traffic Index with Running Mean



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Smoothing, Filtering: Part 2

In the presence a seasonal effect, smoothing approaches are still valid for estimating the trend. We have to make sure that the sum is taken over an entire season, i.e. for monthly data:

$$\hat{m}_t = \frac{1}{12} \left(\frac{1}{2} X_{t-6} + X_{t-5} + \dots + X_{t+5} + \frac{1}{2} X_{t+6} \right) \text{ for } t = 7, \dots, n-6$$

An estimate of the seasonal effect s_t at time t can be obtained by:

$$\hat{s}_t = x_t - \hat{m}_t$$

By averaging these estimates of the effects for each month, we obtain a single estimate of the effect for each month.

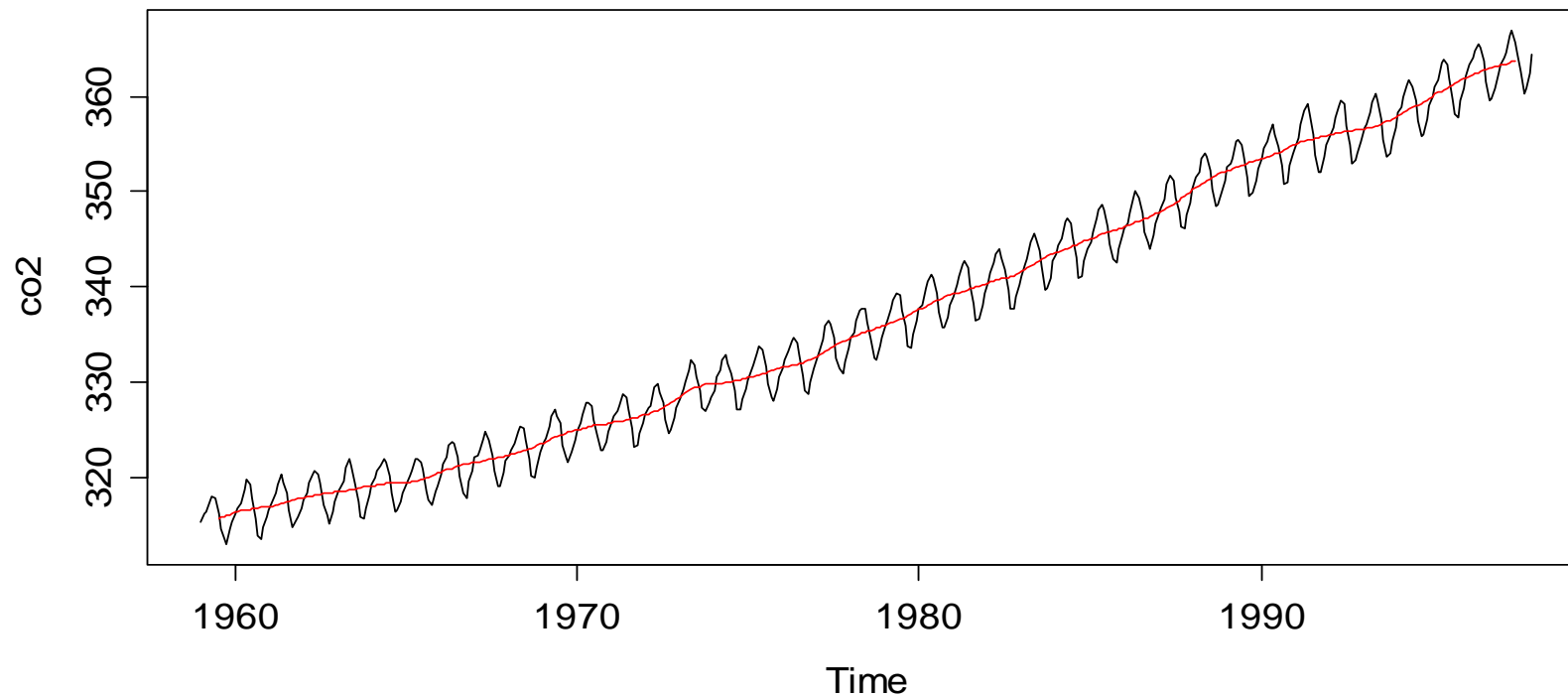
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Trend Estimation for Mauna Loa Data

```
> wghts <- c(.5,rep(1,11),.5)/12  
> trd <- filter(co2, filter=wghts, sides=2)
```

Mauna Loa CO2 Concentrations



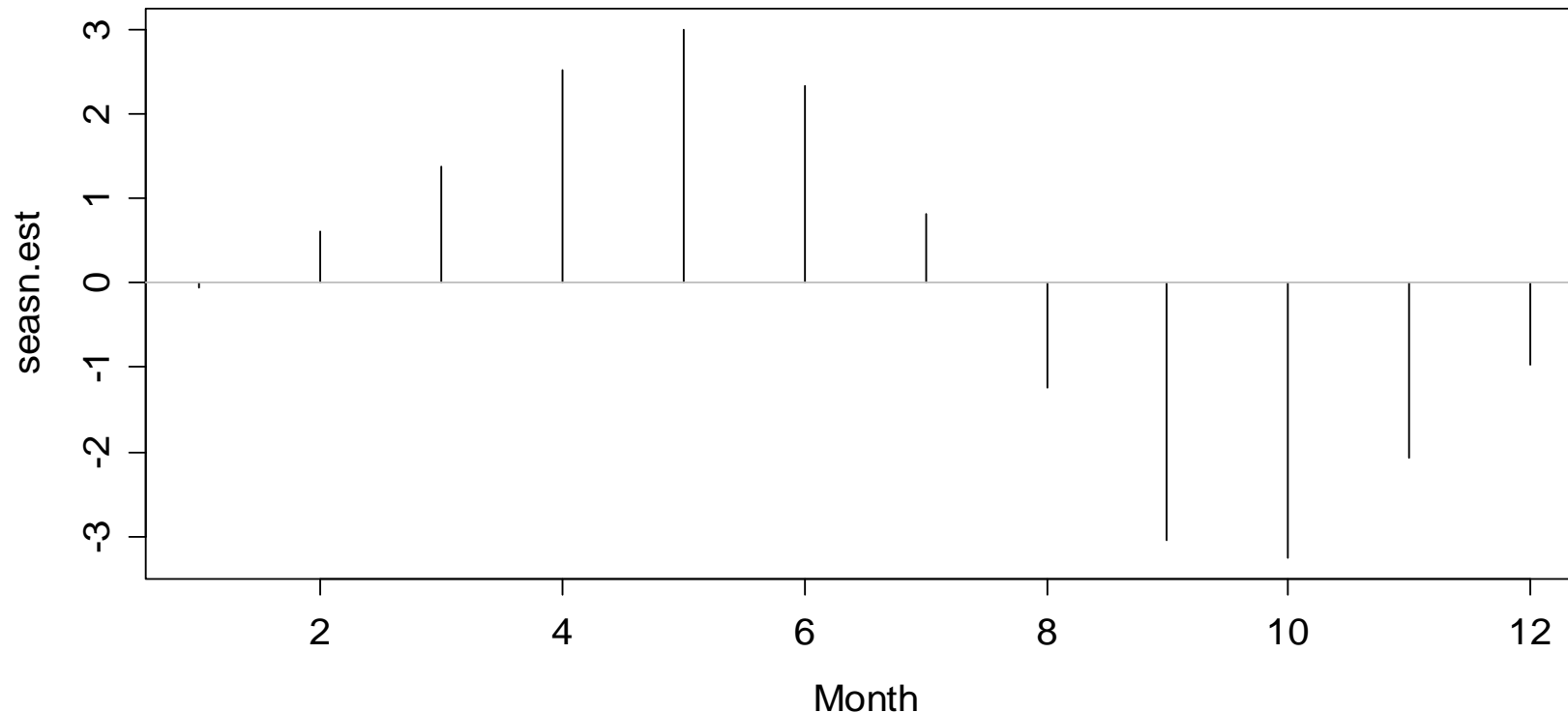
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Estimating the Seasonal Effects

$$\hat{s}_{Jan} = \hat{s}_1 = \hat{s}_{13} = \dots = \frac{1}{39} \cdot \sum_{j=0}^{38} (x_{12j+1} - \hat{m}_{12j+1})$$

Seasonal Effects for Mauna Loa Data



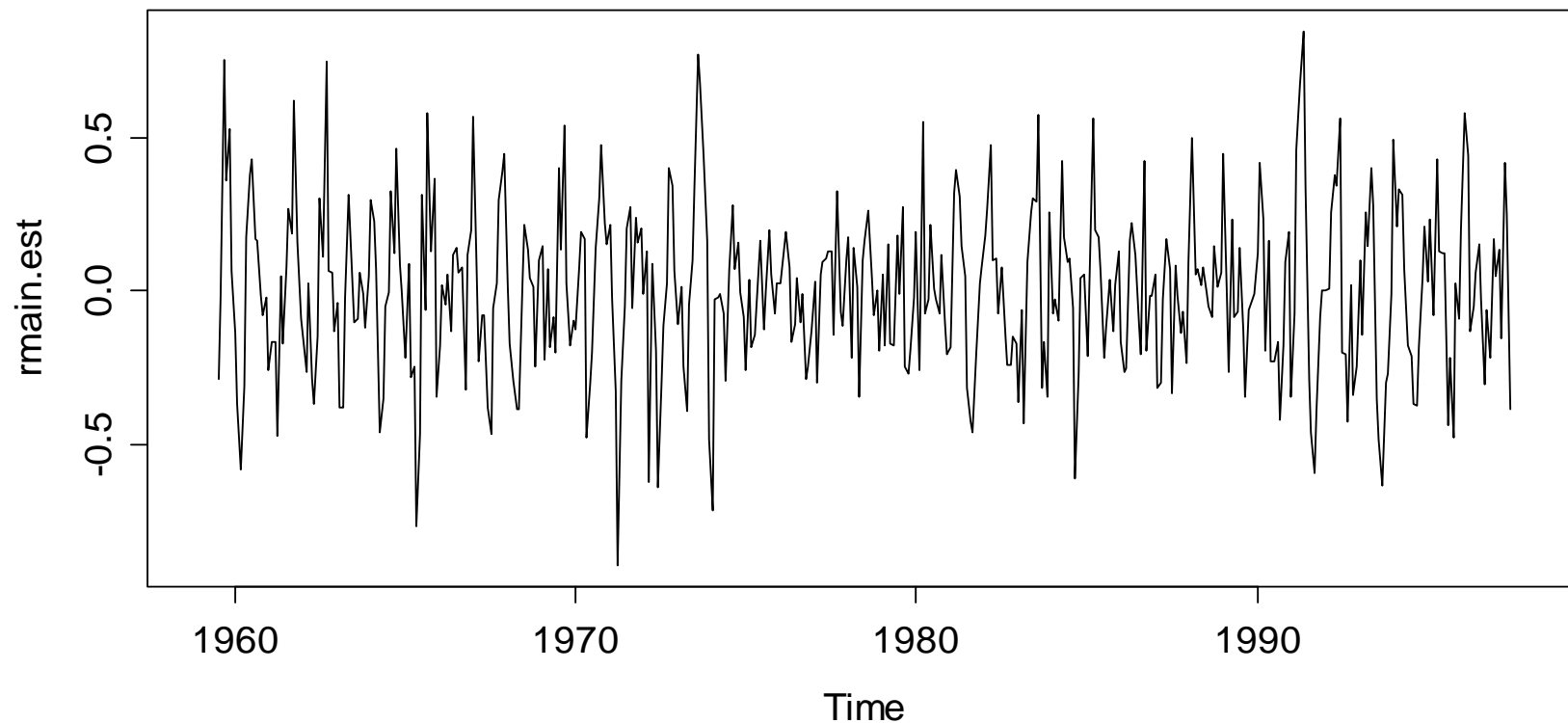
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Estimating the Remainder Term

$$\hat{R}_t = x_t - \hat{m}_t - \hat{s}_t$$

Estimated Stochastic Remainder Term



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Smoothing, Filtering: Part 3

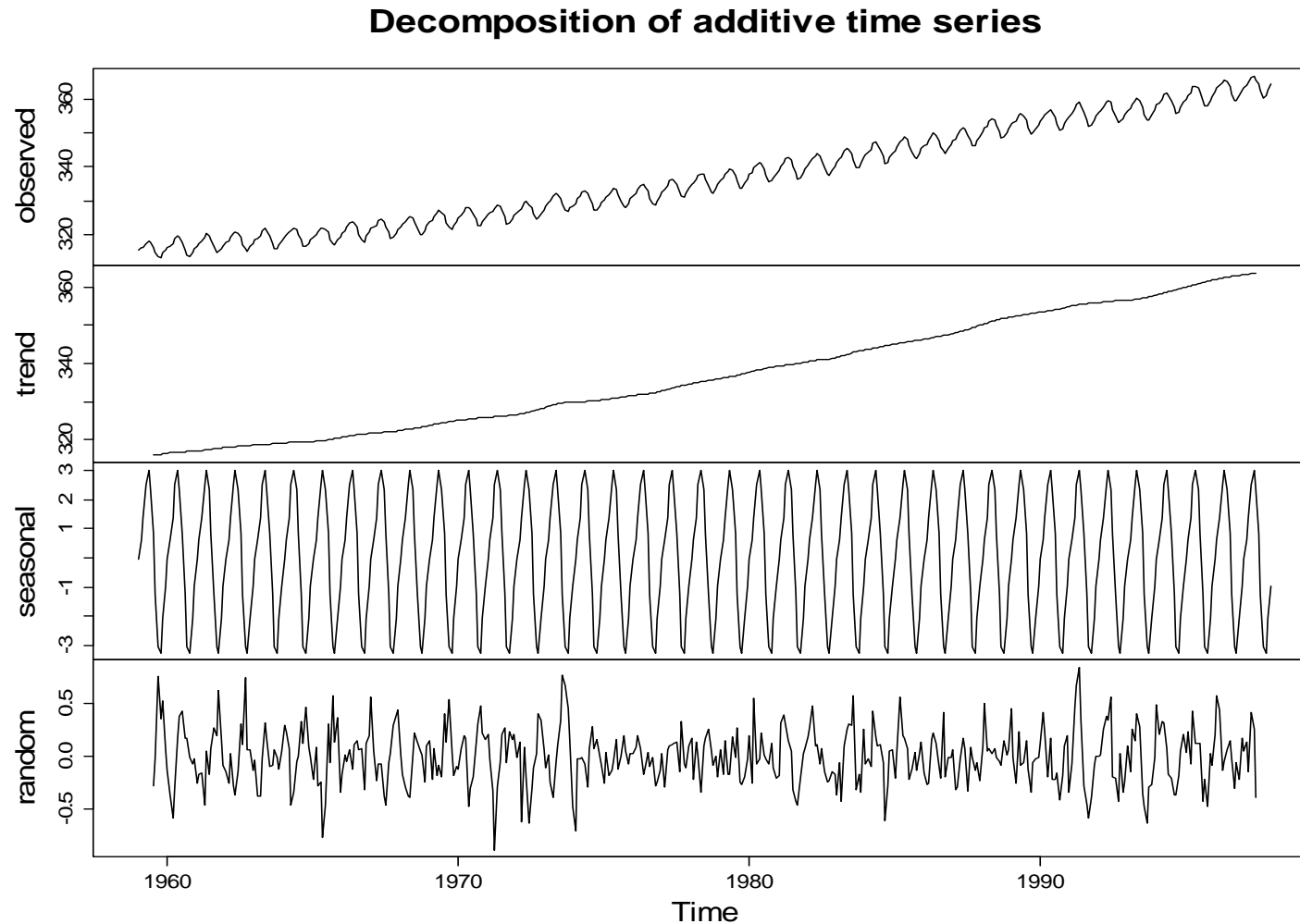
- The smoothing approach is based on estimating the trend first, and then the seasonality.
- The generalization to other periods than $p = 12$, i.e. monthly data is straightforward. Just choose a symmetric window and use uniformly distributed coefficients that sum up to 1.
- The sum over all seasonal effects will be close to zero. Usually, it is centered to be exactly there.
- This procedure is implemented in R with function:
`decompose ()`

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Estimating the Remainder Term

```
> plot(decompose(co2))
```



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Smoothing, Filtering: Remarks

Some advantages and disadvantages:

- + trend and seasonal effect can be estimated
 - + \hat{m}_t , \hat{s}_t and \hat{R}_t are explicitly known, can be visualised
 - + procedure is transparent, and simple to implement
-
- resulting time series will be shorter than the original
 - the running mean is not the very best smoother
 - extrapolation of \hat{m}_t , \hat{s}_t are not entirely obvious